Class-specific Poisson Image Denoising using Importance Sampling

Milad Niknejad, José M. Bioucas-Dias, Mário A. T. Figueiredo

Instituto de Telecomunicações, Instituto Superior Técnico, University of Lisbon, Portugal





Niknejad, Bioucas-Dias, Figueiredo (IST)

- 1 Class-specific image denoising
- 2 Patch estimation using Monte-Carlo
- Importance Sampling
- Applying importance sampling for image denoising
- 5 Proposed method

Often, the image to be denoised belongs to a known specific class,

Examples: text/document, face, fingerprint, a specific type of medical image (*e.g.*, brain MRI), ...

Often, the image to be denoised belongs to a known specific class,

Examples: text/document, face, fingerprint, a specific type of medical image (*e.g.*, brain MRI), ...



Often, the image to be denoised belongs to a known specific class,

Examples: text/document, face, fingerprint, a specific type of medical image (*e.g.*, brain MRI), ...



This knowledge should be exploited by the denoising method!

Class-specific image denoising

Assumption: A dataset of clean images of the same class is available.



Gaussian noise observation model:

$$\mathbf{y}_i = \mathbf{x}_i + \mathbf{v}_i$$

 \mathbf{x}_i is a patch of the original image; \mathbf{y}_i is the corresponding noisy patch; \mathbf{v}_i is i.i.d. Gaussian noise.

Gaussian noise observation model:

$$\mathbf{y}_i = \mathbf{x}_i + \mathbf{v}_i$$

 \mathbf{x}_i is a patch of the original image; \mathbf{y}_i is the corresponding noisy patch; \mathbf{v}_i is i.i.d. Gaussian noise.

Poisson noise observation model (the focus of this presentation):

$$\mathbf{y}_{i,j} \sim \mathcal{P}(\mathbf{x}_{i,j}).$$

 $\mathbf{x}_{i,j}$ is the j^{th} pixel of \mathbf{x}_i . \mathcal{P} is a Poisson distribution with mean $\mathbf{x}_{i,j}$. Gaussian noise observation model:

$$\mathbf{y}_i = \mathbf{x}_i + \mathbf{v}_i$$

 \mathbf{x}_i is a patch of the original image; \mathbf{y}_i is the corresponding noisy patch; \mathbf{v}_i is i.i.d. Gaussian noise.

Poisson noise observation model (the focus of this presentation):

$$\mathbf{y}_{i,j} \sim \mathcal{P}(\mathbf{x}_{i,j}).$$

 $\mathbf{x}_{i,j}$ is the j^{th} pixel of \mathbf{x}_i . \mathcal{P} is a Poisson distribution with mean $\mathbf{x}_{i,j}$.

Goal: recover the clean patch \mathbf{x}_i from the noisy one \mathbf{y}_i

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

• This multi-dimensional integral is intractable, in general (exception: Gaussian noise and Gaussian prior).

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

- This multi-dimensional integral is intractable, in general (exception: Gaussian noise and Gaussian prior).
- Monte-Carlo approximation: obtain samples \mathbf{x}_j from $p(\mathbf{x}|\mathbf{y})$

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

- This multi-dimensional integral is intractable, in general (exception: Gaussian noise and Gaussian prior).
- Monte-Carlo approximation: obtain samples \mathbf{x}_j from $p(\mathbf{x}|\mathbf{y})$

$$\hat{\mathbf{x}}_i = \frac{1}{n} \sum_{j=1}^n \mathbf{x}_j \qquad \lim_{n \to \infty} \hat{\mathbf{x}}_i = \hat{\mathbf{x}}_i$$

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

- This multi-dimensional integral is intractable, in general (exception: Gaussian noise and Gaussian prior).
- Monte-Carlo approximation: obtain samples \mathbf{x}_j from $p(\mathbf{x}|\mathbf{y})$

$$\hat{\mathbf{x}}_i = \frac{1}{n} \sum_{j=1}^n \mathbf{x}_j \qquad \lim_{n \to \infty} \hat{\mathbf{x}}_i = \hat{\mathbf{x}}_i$$

• However, sampling from $p(\mathbf{x}|\mathbf{y}_i)$ is also intractable.

MMSE patch estimate ($p(\mathbf{y} = \mathbf{y}_i)$ is replaced by $p(\mathbf{y}_i)$):

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x} = \int_{\mathbb{R}^p} \mathbf{x} \ \frac{p(\mathbf{y}_i|\mathbf{x}) \ p(\mathbf{x})}{p(\mathbf{y}_i)} \ d\mathbf{x}$$

- This multi-dimensional integral is intractable, in general (exception: Gaussian noise and Gaussian prior).
- Monte-Carlo approximation: obtain samples \mathbf{x}_j from $p(\mathbf{x}|\mathbf{y})$

$$\hat{\mathbf{x}}_i = \frac{1}{n} \sum_{j=1}^n \mathbf{x}_j \qquad \lim_{n \to \infty} \hat{\mathbf{x}}_i = \hat{\mathbf{x}}_i$$

- However, sampling from $p(\mathbf{x}|\mathbf{y}_i)$ is also intractable.
- Can we approximate $\hat{\mathbf{x}}_i$ by sampling from another distribution?

Goal: to compute (or approximate)

$$\mathbb{E}[f(\mathbf{z})] = \int f(\mathbf{z}) \, p(\mathbf{z}) \, d\mathbf{z}.$$

Goal: to compute (or approximate)

$$\mathbb{E}[f(\mathbf{z})] = \int f(\mathbf{z}) \, p(\mathbf{z}) \, d\mathbf{z}.$$

• Let $\tilde{p}(\mathbf{z}) = c p(\mathbf{z})$ be an un-normalized version of $p(\mathbf{z})$.

Goal: to compute (or approximate)

$$\mathbb{E}[f(\mathbf{z})] = \int f(\mathbf{z}) \, p(\mathbf{z}) \, d\mathbf{z}.$$

- Let $\tilde{p}(\mathbf{z}) = c p(\mathbf{z})$ be an un-normalized version of $p(\mathbf{z})$.
- Let q̃(z) = b q(z) be another un-normalized density; assume it is possible/easy to obtain samples z₁,..., z_n ~ q(z).
- Constants c and b may be unknown.

Goal: to compute (or approximate)

$$\mathbb{E}[f(\mathbf{z})] = \int f(\mathbf{z}) \, p(\mathbf{z}) \, d\mathbf{z}.$$

- Let $\tilde{p}(\mathbf{z}) = c p(\mathbf{z})$ be an un-normalized version of $p(\mathbf{z})$.
- Let q̃(z) = b q(z) be another un-normalized density; assume it is possible/easy to obtain samples z₁,..., z_n ~ q(z).
- Constants c and b may be unknown.

$$\hat{\mathbb{E}}_n[f(\mathbf{z})] = \frac{\sum_{j=1}^n f(\mathbf{z}_j) w(\mathbf{z}_j)}{\sum_{j=1}^n w(\mathbf{z}_j)}, \qquad w(\mathbf{z}_j) = \frac{\tilde{p}(\mathbf{z}_j)}{\tilde{q}(\mathbf{z}_j)}.$$

Goal: to compute (or approximate)

$$\mathbb{E}[f(\mathbf{z})] = \int f(\mathbf{z}) \, p(\mathbf{z}) \, d\mathbf{z}.$$

- Let $\tilde{p}(\mathbf{z}) = c p(\mathbf{z})$ be an un-normalized version of $p(\mathbf{z})$.
- Let q̃(z) = b q(z) be another un-normalized density; assume it is possible/easy to obtain samples z₁,..., z_n ~ q(z).
- Constants c and b may be unknown.

$$\hat{\mathbb{E}}_n[f(\mathbf{z})] = \frac{\sum_{j=1}^n f(\mathbf{z}_j) w(\mathbf{z}_j)}{\sum_{j=1}^n w(\mathbf{z}_j)}, \qquad w(\mathbf{z}_j) = \frac{\tilde{p}(\mathbf{z}_j)}{\tilde{q}(\mathbf{z}_j)}.$$

• As in plain Monte-Carlo: $\lim_{n\to\infty} \hat{\mathbb{E}}_n[f(\mathbf{z})] = \mathbb{E}[f(\mathbf{z})]$

Back to our problem:

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x}.$$

Back to our problem:

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x}.$$

• Instead of sampling from $p(\mathbf{x}|\mathbf{y}_i)$, use samples $\mathbf{x}_1, ..., \mathbf{x}_n$ from $p(\mathbf{x})$;

Back to our problem:

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x}.$$

- Instead of sampling from $p(\mathbf{x}|\mathbf{y}_i)$, use samples $\mathbf{x}_1, ..., \mathbf{x}_n$ from $p(\mathbf{x})$;
- Simply use samples from the external dataset of clean patches.

Back to our problem:

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x}.$$

- Instead of sampling from $p(\mathbf{x}|\mathbf{y}_i)$, use samples $\mathbf{x}_1, ..., \mathbf{x}_n$ from $p(\mathbf{x})$;
- Simply use samples from the external dataset of clean patches.
- Use these samples in SNIS

$$\hat{\mathbf{x}}_i = \hat{\mathbb{E}}_n[\mathbf{x}|\mathbf{y}_i] = \frac{\sum_{j=1}^n \mathbf{x}_j w_j}{\sum_{j=1}^n w_j}, \qquad w_j = p(\mathbf{y}_i|\mathbf{x} = \mathbf{x}_j)$$

Back to our problem:

$$\hat{\mathbf{x}}_i = \mathbb{E}[\mathbf{x}|\mathbf{y}_i] = \int_{\mathbb{R}^p} \mathbf{x} \ p(\mathbf{x}|\mathbf{y}_i) \ d\mathbf{x}.$$

- Instead of sampling from $p(\mathbf{x}|\mathbf{y}_i)$, use samples $\mathbf{x}_1, ..., \mathbf{x}_n$ from $p(\mathbf{x})$;
- Simply use samples from the external dataset of clean patches.
- Use these samples in SNIS

$$\hat{\mathbf{x}}_i = \hat{\mathbb{E}}_n[\mathbf{x}|\mathbf{y}_i] = \frac{\sum_{j=1}^n \mathbf{x}_j w_j}{\sum_{j=1}^n w_j}, \qquad w_j = p(\mathbf{y}_i|\mathbf{x} = \mathbf{x}_j)$$

• Why? $\tilde{p}(\mathbf{z}) = p(\mathbf{y}_i | \mathbf{x}) p(\mathbf{x}), c = 1/p(\mathbf{y}_i) \text{ and } \tilde{q}(\mathbf{z}) = p(\mathbf{x}).$

• For Poisson noise, the weights are easy to obtain $(\mathbf{y}_{i,j} \sim \mathcal{P}(\mathbf{x}_{i,j}), \text{ i.i.d.})$

$$w_j = \prod_{l=1}^{N} rac{e^{-\mathbf{x}_{(j,l)}} (\mathbf{x}_{(j,l)})^{\mathbf{y}_{(j,l)}}}{\mathbf{y}_{(j,l)}!}$$

.

• For Poisson noise, the weights are easy to obtain $(\mathbf{y}_{i,j} \sim \mathcal{P}(\mathbf{x}_{i,j}), \text{ i.i.d.})$

$$w_j = \prod_{l=1}^{N} rac{e^{-\mathbf{x}_{(j,l)}} (\mathbf{x}_{(j,l)})^{\mathbf{y}_{(j,l)}}}{\mathbf{y}_{(j,l)}!}$$

It can be adapted to other image restoration tasks, such as deblurring,

$$w_{j} = \prod_{l=1}^{N} \frac{e^{-\mathbf{H}_{(j,l)}\mathbf{x}_{(j,l)}} (\mathbf{H}_{(j,l)}\mathbf{x}_{(j,l)})^{\mathbf{y}_{(j,l)}}}{\mathbf{y}_{(j,l)}!}$$

• For Poisson noise, the weights are easy to obtain $(\mathbf{y}_{i,j} \sim \mathcal{P}(\mathbf{x}_{i,j}), \text{ i.i.d.})$

$$w_j = \prod_{l=1}^{N} rac{e^{-\mathbf{x}_{(j,l)}} (\mathbf{x}_{(j,l)})^{\mathbf{y}_{(j,l)}}}{\mathbf{y}_{(j,l)}!}$$

It can be adapted to other image restoration tasks, such as deblurring,

$$w_{j} = \prod_{l=1}^{N} \frac{e^{-\mathbf{H}_{(j,l)}\mathbf{x}_{(j,l)}} (\mathbf{H}_{(j,l)}\mathbf{x}_{(j,l)})^{\mathbf{y}_{(j,l)}}}{\mathbf{y}_{(j,l)}!}$$

• It can be generalized to other noise models.

Key observations:

 Using samples from p(x) is sub-optimal, as it may have high variance (or even infinite variance). It requires very large n.

Key observations:

- Using samples from p(x) is sub-optimal, as it may have high variance (or even infinite variance). It requires very large n.
- The proposal distribution should be made similar to each target distribution p(x|y_i): Estimator with lower MMSE for limited number of samples.

Key observations:

- Using samples from p(x) is sub-optimal, as it may have high variance (or even infinite variance). It requires very large n.
- The proposal distribution should be made similar to each target distribution p(x|y_i): Estimator with lower MMSE for limited number of samples.

Proposed approach:

Cluster the patches in the external dataset.

Key observations:

- Using samples from p(x) is sub-optimal, as it may have high variance (or even infinite variance). It requires very large n.
- The proposal distribution should be made similar to each target distribution p(x|y_i): Estimator with lower MMSE for limited number of samples.

Proposed approach:

- Cluster the patches in the external dataset.
- Assign each noisy patch to the closest cluster.

Key observations:

- Using samples from p(x) is sub-optimal, as it may have high variance (or even infinite variance). It requires very large n.
- The proposal distribution should be made similar to each target distribution p(x|y_i): Estimator with lower MMSE for limited number of samples.

Proposed approach:

- Cluster the patches in the external dataset.
- Assign each noisy patch to the closest cluster.
- Use the corresponding clean patches as samples from the proposal distribution for SNIS.

Clustering: Any clustering algorithm can be used (k-means,...).
 The whole dataset of patches is clustered to K clusters. {X₁...X_K}.

- Clustering: Any clustering algorithm can be used (k-means,...).
 The whole dataset of patches is clustered to K clusters. {X₁...X_K}.
- Objective is to solve the following simultaneous classification and estimation problem:

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \operatorname*{arg\,min}_{(\mathbf{u},k)} \int_{\mathbb{R}^m_+} \|\mathbf{u} - \mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i, k) \ d\mathbf{x}$$

- Clustering: Any clustering algorithm can be used (k-means,...).
 The whole dataset of patches is clustered to K clusters. {X₁...X_K}.
- Objective is to solve the following simultaneous classification and estimation problem:

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \operatorname*{arg\,min}_{(\mathbf{u},k)} \int_{\mathbb{R}^m_+} \|\mathbf{u} - \mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i,k) \ d\mathbf{x}$$

• The above chooses the best cluster \hat{k}_i , and use this distribution to approximate the integral.

- Clustering: Any clustering algorithm can be used (k-means,...).
 The whole dataset of patches is clustered to K clusters. {X₁...X_K}.
- Objective is to solve the following simultaneous classification and estimation problem:

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \operatorname*{arg\,min}_{(\mathbf{u},k)} \int_{\mathbb{R}^m_+} \|\mathbf{u} - \mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i,k) \ d\mathbf{x}$$

- The above chooses the best cluster \hat{k}_i , and use this distribution to approximate the integral.
- It is equivalent to sampling from (unkown) \hat{k}_i^{th} distribution as the proposal distribution.

- Clustering: Any clustering algorithm can be used (k-means,...).
 The whole dataset of patches is clustered to K clusters. {X₁...X_K}.
- Objective is to solve the following simultaneous classification and estimation problem:

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \operatorname*{arg\,min}_{(\mathbf{u},k)} \int_{\mathbb{R}^m_+} \|\mathbf{u} - \mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i,k) \ d\mathbf{x}$$

- The above chooses the best cluster \hat{k}_i , and use this distribution to approximate the integral.
- It is equivalent to sampling from (unkown) \hat{k}_i^{th} distribution as the proposal distribution.
- The above integral is intractable, but we can use SNIS.

$$\mathbb{E}[\|\mathbf{x}-\mathbf{u}\|_2^2|\mathbf{y}_i,k] = \int_{\mathbb{R}^m_+} \|\mathbf{u}-\mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i,k) \ d\mathbf{x}.$$

2

・ロト ・ 日 ト ・ ヨ ト ・ ヨ ト

$$\mathbb{E}[\|\mathbf{x}-\mathbf{u}\|_2^2|\mathbf{y}_i,k] = \int_{\mathbb{R}^m_+} \|\mathbf{u}-\mathbf{x}\|_2^2 \ p(\mathbf{x}|\mathbf{y}_i,k) \ d\mathbf{x}.$$

Using SNIS, the above can be approximated by

$$\hat{\mathbb{E}}_{n}[\|\mathbf{x} - \mathbf{u}\|_{2}^{2} | \mathbf{y}_{i}, k] = \frac{\sum_{j=1}^{n} \|\mathbf{u} - \mathbf{x}_{j}\|_{2}^{2} p(\mathbf{y}_{i} | \mathbf{x}_{j_{k}})}{\sum_{j=1}^{n} p(\mathbf{y}_{i} | \mathbf{x}_{j_{k}})}$$
(1)

where the \mathbf{x}_{j_k} , for j = 1, ..., n are samples from the distribution $p(\mathbf{x}|k)$.

We minimize the approximation (1) by alternating minimization.

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \underset{(\mathbf{u},k)}{\arg\min} \hat{\mathbb{E}}_n[\|\mathbf{x} - \mathbf{u}\|_2^2 | \mathbf{y}_i, k]$$

3

A B M A B M

Image: Image:

We minimize the approximation (1) by alternating minimization. $(\hat{\mathbf{x}}_i, \hat{k}_i) = \arg\min_{\substack{(\mathbf{u}, k)}} \hat{\mathbb{E}}_n[\|\mathbf{x} - \mathbf{u}\|_2^2 | \mathbf{y}_i, k]$

• when **u** is fixed,

$$\hat{k}_i(\mathbf{u}) = rgmin_k rac{\displaystyle\sum_{j=1}^{n_2} \|\mathbf{u}-\mathbf{x}_j\|_2^2 \ p(\mathbf{y}_i|\mathbf{x}_{j_k})}{\displaystyle\sum_{j=1}^{n_2} p(\mathbf{y}_i|\mathbf{x}_{j_k})}.$$

(2)

We minimize the approximation (1) by alternating minimization.

$$(\hat{\mathbf{x}}_i, \hat{k}_i) = \operatorname*{arg\,min}_{(\mathbf{u},k)} \hat{\mathbb{E}}_n[\|\mathbf{x} - \mathbf{u}\|_2^2 | \mathbf{y}_i, k]$$

when u is fixed,

$$\hat{k}_i(\mathbf{u}) = rgmin_k rac{\displaystyle\sum_{j=1}^{n_2} \|\mathbf{u} - \mathbf{x}_j\|_2^2 \ p(\mathbf{y}_i | \mathbf{x}_{j_k})}{\displaystyle\sum_{j=1}^{n_2} p(\mathbf{y}_i | \mathbf{x}_{j_k})}.$$

• when k is fixed,

$$\hat{\mathbf{x}}_{i}(k) = \hat{\mathbb{E}}_{n_{1}}[\mathbf{x}|\mathbf{y}_{i}, k] = \frac{\sum_{j=1}^{n_{1}} \mathbf{x}_{j} p(\mathbf{y}_{i}|\mathbf{x}_{j})}{\sum_{i=1}^{n_{1}} P(\mathbf{y}_{i}|\mathbf{x}_{j})}$$
(3)

(2)

3

• The key to speeding up is to limit the numbers of patch samples n_1 and n_2 .

- The key to speeding up is to limit the numbers of patch samples n_1 and n_2 .
- Clustering: $n_2 = 30$, overall 600 patches for all k = 20 clusters (less than 1 percent of samples in external datasets).

- The key to speeding up is to limit the numbers of patch samples n_1 and n_2 .
- Clustering: $n_2 = 30$, overall 600 patches for all k = 20 clusters (less than 1 percent of samples in external datasets).
- Denoising: samples derived for each patch n_1 was set to 300.

- The key to speeding up is to limit the numbers of patch samples n_1 and n_2 .
- Clustering: $n_2 = 30$, overall 600 patches for all k = 20 clusters (less than 1 percent of samples in external datasets).
- Denoising: samples derived for each patch n_1 was set to 300.
- Overall: 900 patches are processed for each denoised patch (computational complexity is similar to an internal non-local denoising with the patches constrained in 30×30 window).



ICIP 2017, Beijing, China 15 / 19



Noisy
(Peak=2)

Non-local PCA VST+BM3D Proposed (PSNR=14.95) (PSNR=14.55) (PSNR=18.64) • We proposed a method based on importance sampling in which no parametric distribution is fitted to data

- We proposed a method based on importance sampling in which no parametric distribution is fitted to data
- Any clustering method can be used

- We proposed a method based on importance sampling in which no parametric distribution is fitted to data
- Any clustering method can be used
- Each cluster can be seen as samples of unknown proposal distribution

- We proposed a method based on importance sampling in which no parametric distribution is fitted to data
- Any clustering method can be used
- Each cluster can be seen as samples of unknown proposal distribution
- The method can be generalized easily to other image restoration inverse problems.

A. Owen

Monte Carlo Theory, Methods and Examples Available at http://statweb.stanford.edu/ owen/mc/.

Bugallo, M.F., Martino, L. and Corander, J. Adaptive importance sampling in signal processing Digital Signal Processing 47 (2015): 36-49

Salmon, J., Harmany, Z., Deledalle, C. A., Willett, R. Poisson noise reduction with non-local PCA. Journal of mathematical imaging and vision, 48(2), 279-294.

Thanks for your attention