



Clock synchronization over networks using sawtooth models

Pol del Aguila Pla, Ph.D. https://poldap.github.io, https://github.com/poldap

Biomedical Imaging Group, École Polytechnique Fédérale de Lausanne (EPFL) Mathematical Imaging Section, Center for Biomedical Imaging (CIBM), Switzerland

ICASSP 2020, May 7, 2020

https://github.com/poldap/clock\_sync\_and\_range

Work done at the KTH Royal Institute of Technology

## Clock synchronization





- Introduction to the sawtooth model
- Cramér-Rao Lower Bounds

# Authors and funding



J. Jaldén

L. Pellaco

S. Dwivedi

P. Händel

#### SRA ICT TNG: PITA

T. Oechtering, A. Proutiere, J. Jaldén







Initially proposed in (De Angelis, Dwivedi, Händel, 2013).

Low communication overhead





- Low communication overhead
- Low power consumption





- Low communication overhead
- Low power consumption
- High measurement accuracy





- Low communication overhead
- Low power consumption
- High measurement accuracy
- Nodes can measure their own clock period



















#### Introduction to the sawtooth model (III)

$$y_{\text{det}}[n] = \delta_{\leftrightarrow} + \delta_0 + T_{\mathcal{S}} \left( 1 - \text{mod}_1 \left[ T_{\text{s}} f_{\text{d}} n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} \right] \right)$$



#### Introduction to the sawtooth model (III)

 $Y[n] = \alpha + W[n] + \psi \operatorname{mod}_1(\beta n + \gamma + V[n])$ , with W[n] and V[n] AWGN.



#### Introduction to the sawtooth model (III)

# $Y[n] = \alpha + W[n] + \psi \mod_1(\beta n + \gamma + V[n])$ , with W[n] and V[n] AWGN.



Model derivation

- Estimation theoretic analysis
- Conditions for identifiability
- Exhaustive empirical results
- DOI: 10.1109/OJSP.2020.2978762

Our model is not differentiable, so one can not define the CRLBs.

$$Y[n] = \delta_{\leftrightarrow} + \delta_0 + W[n] + T_{\mathcal{S}} \left[ 1 - \operatorname{mod}_1 \left( T_{\mathrm{s}} f_{\mathrm{d}} n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} + V[n] \right) \right]$$

Our model is not differentiable, so one can not define the CRLBs.

$$Y[n] = \delta_{\leftrightarrow} + \delta_0 + W[n] + T_{\mathcal{S}} \left[ 1 - \operatorname{mod}_{1}(T_{s}f_{d}n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} + V[n]) \right]$$

We work on an unwrapped / linearized model Z[n],

$$Z[n] = \delta_0 + \frac{\delta_{\leftrightarrow}}{2} + T_{\mathcal{S}}\left(1 - \frac{\phi_{\mathcal{S}}}{2\pi}\right) - T_{\mathcal{S}}T_{\rm s}f_{\rm d}n + U[n],$$

with U[n] a white Gaussian process such that  $U[n] \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma^2 = \sigma_w^2 + T_S^2 \sigma_v^2$ .

Our model is not differentiable, so one can not define the CRLBs.

$$Y[n] = \delta_{\leftrightarrow} + \delta_0 + W[n] + T_{\mathcal{S}} \left[ 1 - \operatorname{mod}_{1} (T_{\mathrm{s}} f_{\mathrm{d}} n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} + V[n]) \right]$$

We work on an unwrapped / linearized model Z[n],

$$Z[n] = \delta_0 + \frac{\delta_{\leftrightarrow}}{2} + T_{\mathcal{S}}\left(1 - \frac{\phi_{\mathcal{S}}}{2\pi}\right) - T_{\mathcal{S}}T_{\rm s}f_{\rm d}n + U[n],$$

with U[n] a white Gaussian process such that  $U[n] \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma^2 = \sigma_w^2 + T_S^2 \sigma_v^2$ .

#### Caveats of the approach

Not the same model (but provides a linearized intuition)

Our model is not differentiable, so one can not define the CRLBs.

$$Y[n] = \delta_{\leftrightarrow} + \delta_0 + W[n] + T_{\mathcal{S}} \left[ 1 - \operatorname{mod}_{1} (T_{\mathrm{s}} f_{\mathrm{d}} n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} + V[n]) \right]$$

We work on an unwrapped / linearized model Z[n],

$$Z[n] = \delta_0 + \frac{\delta_{\leftrightarrow}}{2} + T_{\mathcal{S}}\left(1 - \frac{\phi_{\mathcal{S}}}{2\pi}\right) - T_{\mathcal{S}}T_{\rm s}f_{\rm d}n + U[n],$$

with U[n] a white Gaussian process such that  $U[n] \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma^2 = \sigma_w^2 + T_S^2 \sigma_v^2$ .

#### Caveats of the approach

- Not the same model (but provides a linearized intuition)
- Non-identifiability of  $\phi_S$  and  $\delta_{\leftrightarrow}$  (but we can suppose we know the respective other when deriving the lower bound)

Our model is not differentiable, so one can not define the CRLBs.

$$Y[n] = \delta_{\leftrightarrow} + \delta_0 + W[n] + T_{\mathcal{S}} \left[ 1 - \operatorname{mod}_{1} \left( T_{\mathrm{s}} f_{\mathrm{d}} n + \frac{\delta_{\rightarrow}}{T_{\mathcal{S}}} + \frac{\phi_{\mathcal{S}}}{2\pi} + V[n] \right) \right]$$

We work on an unwrapped / linearized model Z[n],

$$Z[n] = \delta_0 + \frac{\delta_{\leftrightarrow}}{2} + T_{\mathcal{S}}\left(1 - \frac{\phi_{\mathcal{S}}}{2\pi}\right) - T_{\mathcal{S}}T_{\rm s}f_{\rm d}n + U[n],$$

with U[n] a white Gaussian process such that  $U[n] \sim \mathcal{N}(0, \sigma^2)$  with  $\sigma^2 = \sigma_w^2 + T_S^2 \sigma_v^2$ .

#### Caveats of the approach

- Not the same model (but provides a linearized intuition)
- Non-identifiability of  $\phi_S$  and  $\delta_{\leftrightarrow}$  (but we can suppose we know the respective other when deriving the lower bound)

• Dependence of 
$$\sigma^2$$
 on  $f_d = \frac{1}{T_S} - \frac{1}{T_M}$  (but that is covered by standard results)

#### Estimation - Periodogram and correlation peaks (PCP)

 $Y[n] = \alpha + W[n] + \psi \operatorname{mod}_1(\beta n + \gamma + V[n])$ , with  $\psi < 0$  known when  $\beta$  is known

Estimate  $|\beta|$  as the highest peak in the periodogram of the L-1-times zero-padded, zero-mean data, i.e.,  $|\hat{\beta}| = \arg \max_k |\text{DFT}_{NL}(\tilde{y}[n])[k]|^2/(NL)$ 



#### Estimation - Periodogram and correlation peaks (PCP)

 $Y[n] = \alpha + W[n] + \psi \operatorname{mod}_1(\beta n + \gamma + V[n])$ , with  $\psi < 0$  known when  $\beta$  is known

- Generate two single-period templates p<sub>±</sub>[n] for ±β and circularly correlate them with the first period of the max-normalized zero-mean data
- Estimate the sign of  $\beta$  by the largest correlation and the phase  $\gamma$  from the index at which it happens, i.e.,  $\hat{\gamma} = \text{mod}_1(\hat{\beta}n^{\text{opt}})$



#### Estimation - Periodogram and correlation peaks (PCP)

 $Y[n] = \alpha + W[n] + \psi \operatorname{mod}_1(\beta n + \gamma + V[n])$ , with  $\psi < 0$  known when  $\beta$  is known

- Estimate  $|\beta|$  as the highest peak in the periodogram of the L-1-times zero-padded, zero-mean data, i.e.,  $|\hat{\beta}| = \arg \max_k |\text{DFT}_{NL}(\tilde{y}[n])[k]|^2/(NL)$
- Generate two single-period templates p<sub>±</sub>[n] for ±β and circularly correlate them with the first period of the max-normalized zero-mean data
- Estimate the sign of  $\beta$  by the largest correlation and the phase  $\gamma$  from the index at which it happens, i.e.,  $\hat{\gamma} = \text{mod}_1(\hat{\beta}n^{\text{opt}})$
- Estimate  $\psi$  through its known relation with  $\beta$ , and  $\alpha$  by the closed-form minimum prediction mean squared error solution assuming  $\hat{\beta}$ ,  $\hat{\psi}$  and  $\hat{\gamma}$  are correct, i.e.,

$$\hat{\alpha}_{\hat{\beta},\hat{\gamma}} = \sum_{n=0}^{N-1} y[n] - \sum_{m=0}^{N-1} \hat{\psi}_{\hat{\beta}} \operatorname{mod}_1 \left[ \hat{\beta} m + \hat{\gamma} \right].$$

# Estimation - Grid Search, either local (LGS), or global (GGS)

Define a grid  $\mathcal{G} \times \mathcal{B}$  in  $\left[-\frac{1}{2}, \frac{1}{2}\right) \times [0, 1)$ , and estimate the point in the grid that minimizes the prediction mean squared error, i.e.,

$$\min_{(\beta,\gamma)\in\mathcal{G}\times\mathcal{B}}\left\{\sum_{n=0}^{N-1}\left(y[n]-\hat{\alpha}_{\beta,\gamma}-\hat{\psi}_{\beta}\operatorname{mod}_{1}[\beta n+\gamma]\right)^{2}\right\},$$

where  $\hat{\alpha}_{\beta,\gamma}$  is the closed form solution as above, and  $\hat{\psi}_{\beta}$  is the known amplitude given the frequency.

# Estimation - Grid Search, either local (LGS), or global (GGS)

Define a grid  $\mathcal{G} \times \mathcal{B}$  in  $\left[-\frac{1}{2}, \frac{1}{2}\right) \times [0, 1)$ , and estimate the point in the grid that minimizes the prediction mean squared error.

 $\blacktriangleright \ \mathcal{G} \times \mathcal{B}$  can be chosen global, resulting on a very irregular function landscape [dB]



# Estimation - Grid Search, either local (LGS), or global (GGS)

Define a grid  $\mathcal{G} \times \mathcal{B}$  in  $\left[-\frac{1}{2}, \frac{1}{2}\right) \times [0, 1)$ , and estimate the point in the grid that minimizes the prediction mean squared error.

- $\blacktriangleright~\mathcal{G}\times\mathcal{B}$  can be chosen global, resulting on a very irregular function landscape
- or local, around the result of PCP, where smoother behavior is expected and better estimates are likely due to a finer gridding if the PCP was close to the right solution



#### Empirical results



#### Empirical results



#### Empirical results







ICASSPE



#### using sawtooth models

Pol del Aguila Pla, Ph.D. https://poldap.github.io, https://github.com/poldap

Biomedical Imaging Group, École Polytechnique Fédérale de Lausanne (EPFL) Mathematical Imaging Section, Center for Biomedical Imaging (CIBM), Switzerland

ICASSP 2020, May 7, 2020 DOI: 10.1109/ICASSP40776.2020.9054426

https://github.com/poldap/clock\_sync\_and\_range

Work done at the KTH Royal Institute of Technology

#### CLOCK STACHBOALLABOA OVER NETWORKS UNIVE SONTOOTH MOR

Remedical Imaging Group, 1795. Lanuaran, Stritowland inian of Information Reimor and Engineering. Advant of HE KTH Reyal Institute of Technology, Stockholm, Tarvins

-------