## Optimal Local Thresholds for Distributed Detection in Energy Harvesting Wireless Sensor Networks

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- Introduction
- System Model and problem statement
- Optimizing local decision thresholds
- Simulation results
- Conclusions



- The designs of wireless sensor networks to perform the task of distributed detection are often based on the conventional battery-powered sensors, leading into designs with a short lifetime, due to battery depletion.
- Energy harvesting, which can collect energy from renewable resources of environment (e.g., solar, wind, and geothermal energy) promises a self-sustainable system with a lifetime.





Figure 1: Our System model during one observation period.



Let  $x_k$  denote the local observation at sensor k:

$$x_k = egin{cases} g_k \mathcal{A} + w_k & \mathcal{H}_1 \ w_k & \mathcal{H}_0 \end{cases}$$

- $\mathcal{A}$  is a known scalar signal
- $w_k \sim \mathcal{N}(0, \sigma_{w_k}^2) \rightarrow \text{Additive noise}$
- $g_k \sim \mathcal{N}(0, \gamma_{g_k}) \rightarrow \mathsf{Multiplicative}$  noise
- All observation noises are independent over time and among *K* sensors.



(1)

During each observation period, sensor k takes N samples of  $x_k$  to measure the received signal energy and applies an energy detector to make a binary decision, i.e., sensor k decides whether or not signal A is present.

$$\Lambda_{k} = \frac{1}{N} \sum_{n=1}^{N} |x_{k,n}|^{2} \gtrsim \frac{d_{k}=1}{d_{k}=0} \ \theta_{k}$$
(2)

• 
$$P_{f_k} = \Pr(\Lambda_k > \theta_k | \mathcal{H}_0) = \frac{\Gamma(N/2, \frac{N\theta_k}{\sigma_{w_k}})}{\Gamma(N/2)}$$
  
•  $P_{d_k} = \Pr(\Lambda_k > \theta_k | \mathcal{H}_1) = Q_{N/2}(\frac{\sqrt{\eta_k}}{\sigma_{w_k}}, \frac{\sqrt{N\theta_k}}{\sigma_{w_k}})$   
• Our goal is optimize the local decision threshold



 $\theta_{k}$ 

Assumptions:

- Each sensor is able to harvest energy from the environment and stores it in a battery with the capacity  $\mathcal{K}$  units of energy.
- The sensors communicate with the FC through orthogonal fading channels with channel gains  $|h_k|$ 's with parameters  $\gamma_{h_k}$ .
- The sensors employ on-off keying signaling.
- We use the channel-inversion power, the number of energy units spent to convey a decision is inversely proportional to  $|h_k|$ .
- To avoid the battery depletion when  $|h_k|$  is too small, we impose an extra constraint for channel quality.



Let  $u_{k,t}$  represent the sensor output corresponding to the observation period t.

$$u_{k,t} = \begin{cases} \lceil \frac{\lambda}{|h_k|} \rceil & \Lambda_k > \theta_k, \ b_{k,t} > \lceil \frac{\lambda}{|h_k|} \rceil, \ |h_k|^2 > \zeta_k \\ 0 & \text{Otherwise} \end{cases}$$
(3)

- $b_{k,t}$  denote the battery state of sensor k
- $|h_k|$  is channel gain
- $\zeta_k$  is threshold of the channel quality
- $\lambda$  is a power regulation constant



We model  $b_{k,t}$  in (3) as the following

$$b_{k,t} = \min\left\{b_{k,t-1} - \lceil \frac{\lambda}{|h_k|} \rceil I_{u_{k,t-1}} + \Omega_{k,t} , \mathcal{K}\right\}$$
(4)

•  $\Omega_{k,t} \in \{0,1\}$  indicates units of harvesting energy and it is a Bernoulli random variable, with  $\Pr(\Omega_{k,t}=1)=p_e$ 

• 
$$I_{u_{k,t-1}} = \begin{cases} 1 & u_{k,t-1} > 0 \\ 0 & \text{Otherwise} \end{cases}$$



#### Battery Model

• Assuming  $b_k$  in (4) is a stationary random process, one can compute the CDF and the pmf of  $b_k$  in terms of  $\mathcal{K}, p_e, \gamma_{h_k}$ . Further, we use pmf of  $b_k$  for our numerical results.



Figure 2: (a) CDF of  $b_k$  for  $\mathcal{K}=20$  and  $p_e=0.5, 0.75, 0.82$ , (b) pmf of  $b_k$  for  $\mathcal{K}=50$  and  $p_e=0.8$ .

We consider two detection performance metrics to find the optimal  $\theta_k$ 's:

• The detection probability at the FC, assuming that the FC utilizes the optimal fusion rule based on Neyman-Pearson optimality criterion.

• the KL distance between the two distributions of the received signals at the FC conditioned on hypothesis  $\mathcal{H}_0, \mathcal{H}_1$ 



The received signal at the FC from sensor k is  $y_k = h_k u_k + n_k$ , where the additive communication channel noise  $n_k \sim \mathcal{N}(0, \sigma_{n_k}^2)$ . The likelihood ratio at the FC is

$$\Delta_{\text{LRT}} = \sum_{k=1}^{K} \log \left( \frac{\sum_{u_k} f(y_k | u_k, \mathcal{H}_1) \operatorname{Pr}(u_k | \mathcal{H}_1)}{\sum_{u_k} f(y_k | u_k, \mathcal{H}_0) \operatorname{Pr}(u_k | \mathcal{H}_0)} \right)$$
(5)

Given  $u_k$ ,  $y_k$  is Gaussian, i.e.,  $y_k|_{u_k=0} \sim \mathcal{N}\left(0, \sigma_{n_k}^2\right)$  and  $y_k|_{u_k=\lceil \frac{\lambda}{|h_k|}\rceil} \sim \mathcal{N}\left(\lceil \frac{\lambda}{|h_k|}\rceil h_k, \sigma_{n_k}^2\right)$ .



### Optimal LRT Fusion Rule and $P_D$ , $P_F$ Expressions

The probabilities  $Pr(u_k|\mathcal{H}_1)$ ,  $Pr(u_k|\mathcal{H}_0)$  in (5) are

• 
$$\Pr\left(u_{k} = \left\lceil \frac{\lambda}{|h_{k}|} \right\rceil | \mathcal{H}_{1}\right) = P_{d_{k}}\rho_{k}q_{k} = \alpha_{k}$$
  
•  $\Pr\left(u_{k} = \left\lceil \frac{\lambda}{|h_{k}|} \right\rceil | \mathcal{H}_{0}\right) = P_{f_{k}}\rho_{k}q_{k} = \beta_{k}$   
where  $\rho_{k} = \Pr(b_{k} > \left\lceil \frac{\lambda}{|h_{k}|} \right\rceil)$  and  $q_{k} = \Pr(|h_{k}|^{2} > \zeta_{k})$ .  
Given a threshold  $\tau$ , the optimal likelihood ratio test (LRT) is  
 $\Delta_{\text{LRT}} \gtrsim \frac{\mathcal{H}_{1}}{\mathcal{H}_{0}}\tau$ . The  $P_{F}, P_{D}$  at the FC  
 $P_{F} = \Pr\left(\Delta_{\text{LRT}} > \tau | \mathcal{H}_{0}\right) = Q\left(\frac{\tau - \mu_{\Delta}|\mathcal{H}_{0}}{\sigma_{\Delta}|\mathcal{H}_{0}}\right)$   
 $P_{D} = \Pr\left(\Delta_{\text{LRT}} > \tau | \mathcal{H}_{1}\right)$   
 $= Q\left(\frac{Q^{-1}(a)\sigma_{\Delta}|\mathcal{H}_{0} + \mu_{\Delta}|\mathcal{H}_{0} - \mu_{\Delta}|\mathcal{H}_{1}}{\sigma_{\Delta}|\mathcal{H}_{1}}\right)$ 



(6)

Kullback-Leibler distance (KL) between the two distributions of the received signals at the FC  $\,$ 

$$KL_{k} = \int_{y_{k}} f(y_{k}|\mathcal{H}_{1}) \log\left(\frac{f(y_{k}|\mathcal{H}_{1})}{f(y_{k}|\mathcal{H}_{0})}\right) dy_{k}$$
(8)

One can approximate  $KL_k$  in (8) by the KL distance of two Gaussian distributions

$$KL_{k} \approx \frac{1}{2} \log(\frac{\sigma_{y_{k}|\mathcal{H}_{0}}^{2}}{\sigma_{y_{k}|\mathcal{H}_{1}}^{2}}) + \frac{\sigma_{y_{k}|\mathcal{H}_{1}}^{2} - \sigma_{y_{k}|\mathcal{H}_{0}}^{2} + (\mu_{y_{k}|\mathcal{H}_{1}} - \mu_{y_{k}|\mathcal{H}_{0}})^{2}}{2\sigma_{y_{k}|\mathcal{H}_{0}}^{2}} \quad (9)$$



In this section, we consider:

- Scheme I: Numerically find  $\theta_k$ 's which maximize  $P_D$  in (7)  $\rightarrow$  *K*-dimensional search is required  $\rightarrow$  computational complexity!
- Scheme II: Finding θ<sub>k</sub>'s which maximize KL<sub>tot</sub> = ∑<sup>K</sup><sub>k=1</sub> KL<sub>k</sub>, using the KL<sub>k</sub> approximation in (9) → Only one dimensional search → computationally efficient.
- Special case: Assume all sensors employ the same local threshold  $\theta_k = \theta$  and compare schemes I and II.

We then compare  $P_D$  evaluated at the  $\theta_k$ 's obtained from mentioned schemes.



### Simulation results



Figure 3: (a)  $P_D$  vs.  $P_F$ (b)  $P_D$  vs.  $\mathcal{K}$ 



- We studied a distributed detection problem in a wireless network with *K* heterogeneous energy harvesting sensors and investigated the optimal local decision thresholds for given transmission and battery state models.
- Our numerical results indicate that the thresholds obtained from maximizing the KL distance are near-optimal and computationally very efficient, as it requires only K one-dimensional searches, as opposed to a K-dimensional search required to find the thresholds that maximize the detection probability.
- The performance gap between each scheme and its corresponding special case indicates that when sensors are heterogeneous, it is advantageous to use different local thresholds according to sensors' statistics.



# Questions?



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