

SPIKE-AND-SLAB VARIATIONAL INFERENCE FOR BLIND IMAGE DECONVOLUTION



UNIVERSIDAD
DE GRANADA

Juan G. Serra¹, Javier Mateos¹, Rafael Molina¹, Aggelos K. Katsaggelos²

Contact email: jmd@decsai.ugr.es

¹Departamento de Ciencias de la Computación e Inteligencia Artificial, Universidad de Granada, Spain.

²Electrical Engineering and Computer Science Department, Northwestern University, Evanston, IL, USA.



Northwestern
University

1. Introduction

Blind image deconvolution (BID) aims at retrieving the original sharp image \mathbf{x} from a blurry and noisy observation \mathbf{y} .

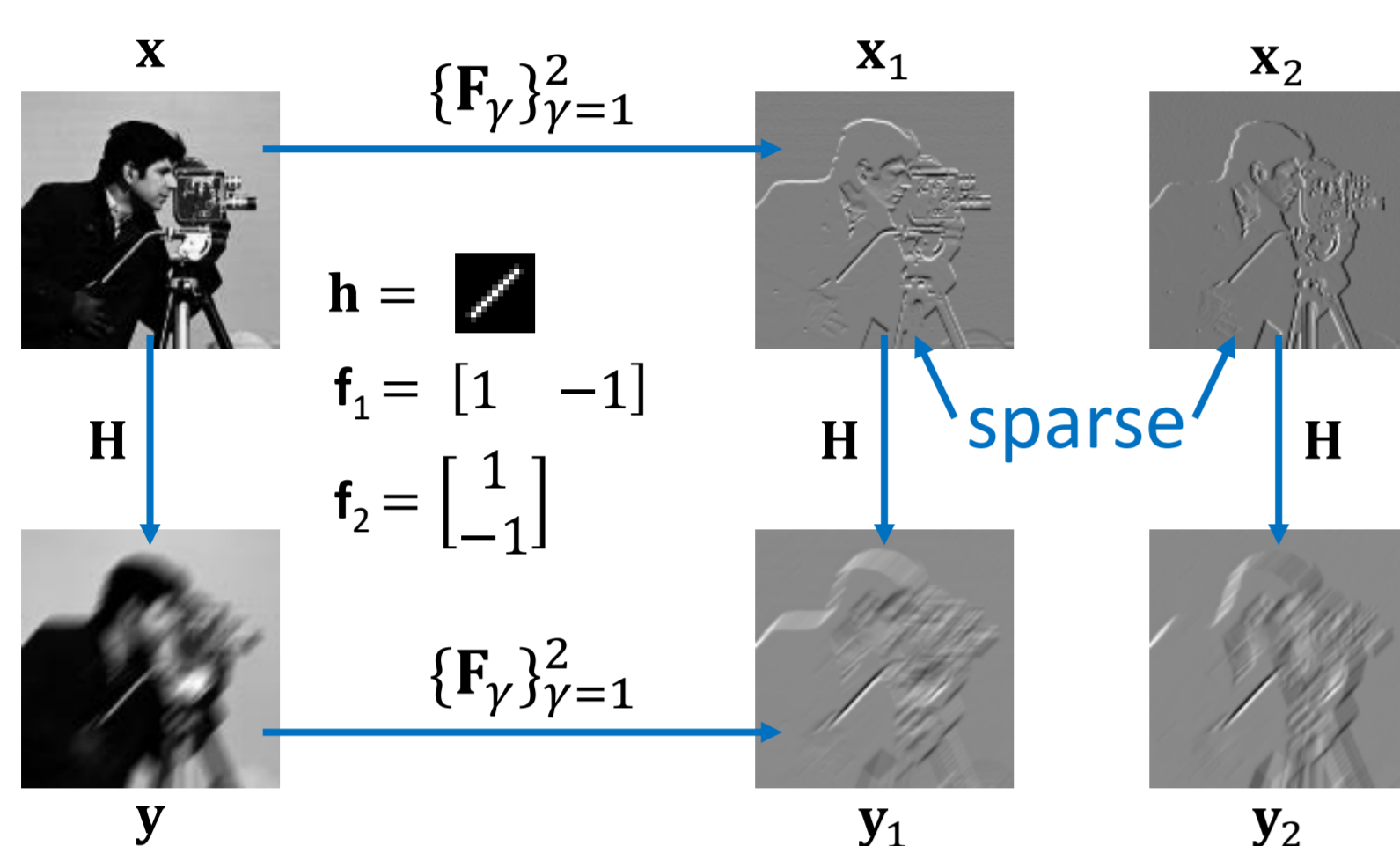
$$\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n} \quad \text{noise}$$

where \mathbf{x} denotes the original image and \mathbf{H} is the convolution matrix associated with the **unknown blur kernel** \mathbf{h} .

We formulate the BID problem in the filter space, creating a set of L pseudo-observations

$$\mathbf{y}_\gamma = \mathbf{F}_\gamma \mathbf{y} = \mathbf{H}\mathbf{F}_\gamma \mathbf{x} + \mathbf{F}_\gamma \mathbf{n} = \mathbf{H}\mathbf{x}_\gamma + \mathbf{n}_\gamma$$

Notice that $\{\mathbf{x}_\gamma\}_{\gamma=1}^L$ should be sparse since they represent high-pass filtered instances of the original image.



We develop a **Blind Image Deconvolution** method that

- uses **Bayesian Modeling and Variational Inference**
 - utilizes the **Spike-and-Slab** prior to impose sparsity
 - is more robust against noise
- The variational inference we propose
- is **more accurate** than the standard mean field variational approximation
 - is much **more efficient** than MCMC

2. Bayesian Modeling

Observation model (in the filter space)

$$p(\mathbf{y}_\Gamma | \mathbf{h}, \mathbf{x}_\Gamma, \beta_\Gamma) = \prod_{\gamma=1}^L p(\mathbf{y}_\gamma | \mathbf{h}, \mathbf{x}_\gamma, \beta_\gamma) \quad \text{noise precision}$$

$$= \prod_{\gamma=1}^L \mathcal{N}(\mathbf{y}_\gamma | \mathbf{H}\mathbf{x}_\gamma, \beta_\gamma^{-1} \mathbf{I})$$

Prior model

We impose a **Spike-and-Slab** prior on each pixel $x_{\gamma i}$

$$p(x_{\gamma i} | \alpha_\gamma, \pi_\gamma) = \pi_\gamma \mathcal{N}(x_{\gamma i} | 0, \alpha_\gamma^{-1}) + (1 - \pi_\gamma) \delta(x_{\gamma i}),$$

Slab: represents the edges of the image

Spike: Dirac delta, represents the flat areas of the image

This is a truly sparse prior: $x_{\gamma i}$ is exactly 0 with probability $1 - \pi_\gamma$.

Unfortunately, variational inference with this prior is intractable.

However, we then rewrite $x_{\gamma i}$ as the product of a Gaussian zero-mean random variable $\tilde{x}_{\gamma i}$ and a Bernoulli random variable $s_{\gamma i}$.

$$x_{\gamma i} = s_{\gamma i} \tilde{x}_{\gamma i}$$

and redefine the prior on the two components of $x_{\gamma i}$ as

$$p(\tilde{x}_{\gamma i}, s_{\gamma i} | \alpha_\gamma, \pi_\gamma) = \mathcal{N}(\tilde{x}_{\gamma i} | 0, \alpha_\gamma^{-1}) \pi_\gamma^{s_{\gamma i}} (1 - \pi_\gamma)^{1 - s_{\gamma i}}$$

where $s_{\gamma i} \in \{0, 1\}$, which is tractable.

Joint probability distribution

With all the above, we have

$$p(\Theta, \mathbf{y}_\Gamma) = p(\mathbf{h}) \prod_{\gamma=1}^L p(\mathbf{y}_\gamma | \mathbf{h}, \tilde{\mathbf{x}}_\gamma, \mathbf{s}_\gamma, \beta_\gamma) \times \prod_{\gamma=1}^L \prod_i p(\tilde{x}_{\gamma i}, s_{\gamma i} | \alpha_\gamma, \pi_\gamma).$$

set of unknowns
 $\Theta = \{\mathbf{h}, \tilde{\mathbf{x}}_\gamma, \mathbf{s}_\gamma\}$

Spike-and-slab-prior

Flat prior on \mathbf{h}

3. Variational Bayesian Inference

Since $p(\Theta | \mathbf{y})$ cannot be calculated in closed form, the standard mean field approximation that factorizes $q(\tilde{\mathbf{x}}_\gamma, \mathbf{s}_\gamma) = q(\tilde{\mathbf{x}}_\gamma)q(\mathbf{s}_\gamma)$ could be used. However this is a unimodal distribution [1] and, therefore, not a good approximation of the true posterior distribution. Since the pairs $\{\tilde{x}_{\gamma i}, s_{\gamma i}\}$ are strongly correlated (recall that $x_{\gamma i} = \tilde{x}_{\gamma i} s_{\gamma i}$), we treat them as a unit, hence we use the factorization

$$q(\mathbf{x}_\gamma, \mathbf{s}_\gamma) = \prod_{\gamma=1}^L \prod_{i=1}^N q(\tilde{x}_{\gamma i}, s_{\gamma i})$$

and utilize the following mean field approximation

$$q(\Theta) = q(\mathbf{h}) \prod_{\gamma=1}^L \prod_{i=1}^N q(\tilde{x}_{\gamma i}, s_{\gamma i}).$$

Obtaining $q(\tilde{\mathbf{x}}_\gamma, \mathbf{s}_\gamma)$

Using the Kullback-Leibler criterion and the mean field factorization presented above, we have

$$q(\tilde{x}_{\gamma i}, s_{\gamma i}) = \frac{1}{Z} \exp \left[\langle \ln p(\mathbf{y}_\gamma | \mathbf{h}, \tilde{\mathbf{x}}_\gamma, \mathbf{s}_\gamma, \beta_\gamma) \rangle_{q(\Theta_{\tilde{x}_{\gamma i}, s_{\gamma i}})} \right] \times \mathcal{N}(\tilde{x}_{\gamma i} | 0, \alpha_\gamma^{-1}) \pi_\gamma^{s_{\gamma i}} (1 - \pi_\gamma)^{1 - s_{\gamma i}}.$$

To compute the explicit expression for the above posterior we separate the derivations for $q(s_{\gamma i})$ and $q(\tilde{x}_{\gamma i} | s_{\gamma i})$.

The distributions $q(s_{\gamma i} = 0)$ and $q(s_{\gamma i} = 1)$ can be obtained by marginalization. Defining

$$\omega_{\gamma i} = q(s_{\gamma i} = 1) = \frac{1}{1 + e^{-u_{\gamma i}}}$$

we have

$$u_{\gamma i} = \ln q(s_{\gamma i} = 1) - \ln q(s_{\gamma i} = 0) = \ln \frac{\pi_\gamma}{1 - \pi_\gamma} + \frac{1}{2} \ln \alpha_\gamma - \frac{1}{2} \ln(\rho_\gamma) + \frac{\beta_\gamma^2}{2\rho_\gamma} (\mathbf{h}_i^\top (\mathbf{y}_\gamma - \sum_{k \neq i} \langle s_{\gamma k} \tilde{x}_{\gamma k} \rangle \mathbf{h}_k))^2.$$

It can be shown that the conditional distributions $q(\tilde{x}_{\gamma i} | s_{\gamma i})$ for $s_{\gamma i} \in \{0, 1\}$ are both Gaussians of the form

$$q(\tilde{x}_{\gamma i} | s_{\gamma i} = 0) = \mathcal{N}(\tilde{x}_{\gamma i} | 0, \alpha_\gamma^{-1}),$$

$$q(\tilde{x}_{\gamma i} | s_{\gamma i} = 1) = \mathcal{N}(\tilde{x}_{\gamma i} | \mu_{x_{\gamma i}}, \rho_\gamma^{-1}),$$

with

$$\mu_{x_{\gamma i}} = \frac{\beta_\gamma}{\rho_\gamma} \mathbf{h}_i^\top (\mathbf{y}_\gamma - \sum_{k \neq i} \langle s_{\gamma k} \tilde{x}_{\gamma k} \rangle \mathbf{h}_k)$$

$$\rho_\gamma = \beta_\gamma \|\mathbf{h}\|^2 + \alpha_\gamma.$$

Finally, we can express the posterior as

$$q(\tilde{x}_{\gamma i}, s_{\gamma i}) = q(\tilde{x}_{\gamma i} | s_{\gamma i}) q(s_{\gamma i}) = \omega_{\gamma i}^{s_{\gamma i}} (1 - \omega_{\gamma i})^{1 - s_{\gamma i}} \times \mathcal{N}(\tilde{x}_{\gamma i} | s_{\gamma i} \mu_{x_{\gamma i}}, s_{\gamma i} \rho_\gamma^{-1} + (1 - s_{\gamma i}) \alpha_\gamma^{-1}).$$

Furthermore,

$$\langle s_{\gamma i} \tilde{x}_{\gamma i} \rangle = \langle x_{\gamma i} \rangle = \omega_{\gamma i} \mu_{x_{\gamma i}},$$

$$\langle s_{\gamma i}^2 \tilde{x}_{\gamma i}^2 \rangle = \langle s_{\gamma i} \tilde{x}_{\gamma i}^2 \rangle = \langle x_{\gamma i}^2 \rangle = \omega_{\gamma i} (\mu_{x_{\gamma i}}^2 + \rho_\gamma^{-1}).$$

Obtaining $q(\mathbf{h})$

Notice that we assume a degenerate distribution on $q(\mathbf{h})$, which leads to the point estimate for \mathbf{h}

$$\hat{\mathbf{h}} = \arg \min_{\mathbf{h}} \sum_{\gamma=1}^L \langle \|\mathbf{y}_\gamma - \mathbf{H}\mathbf{x}_\gamma\|^2 \rangle_{q(\Theta_{\mathbf{h}})}$$

$$= \arg \min_{\mathbf{h}} \sum_{\gamma=1}^L \left[\|\mathbf{y}_\gamma - \mathbf{H} \langle \mathbf{x}_\gamma \rangle\|^2 + \sum_{i=1}^N (\langle x_{\gamma i}^2 \rangle - \langle x_{\gamma i} \rangle^2) \|\mathbf{h}_i\|^2 \right],$$

constrained to $h_j \geq 0$, $\sum_j h_j = 1$.

We can efficiently solve this minimization problem with the ADMM method in [2].

Final image estimation

Once the estimate of the blur, $\hat{\mathbf{h}}$, has been obtained, a non-blind deconvolution algorithm is used to recover an estimation of the original sharp image by solving the problem

$$\hat{\mathbf{x}} = \arg \min_{\mathbf{x}} \frac{1}{2} \|\hat{\mathbf{H}}\mathbf{x} - \mathbf{y}\|_2^2 + \frac{\lambda}{p} \sum_{\gamma=1}^L \|\mathbf{x}_\gamma\|_p,$$

using the fast iterative method in [2].

[1] M. K. Titsias and M. Lázaro-Gredilla, "Spike and slab variational inference for multi-task and multiple kernel learning," in NIPS, 2011, pp. 2339–2347.

[2] Zhou, M. Vega, F. Zhou, R. Molina, and A. K. Katsaggelos, "Fast Bayesian blind deconvolution with Huber super Gaussian priors," Digit. Signal Process., vol. 60, pp. 122–133, 2017.

4. Blind Deconvolution Algorithm

Algorithm 1 Blur estimation algorithm

Require: $\mathbf{y}_\Gamma, \beta_\Gamma, \alpha_\Gamma, \pi_\Gamma$, and initial values for \mathbf{h}^0 , and $\langle \mathbf{x}_\Gamma \rangle^{(0)}$.
Set $k = 0$.
repeat
 Update $\mu_{x_{\gamma i}}^{(k)}, \rho_{\gamma i}^{(k)}$, and $\omega_{\gamma i}^{(k)}, \forall \gamma, \forall i$.
 Update $\langle x_{\gamma i} \rangle$ and $\langle x_{\gamma i}^2 \rangle, \forall \gamma, \forall i$.
 Update $\mathbf{h}^{(k+1)}$.
 Set $k = k + 1$.
until convergence.
return $\hat{\mathbf{h}} = \mathbf{h}^{(k)}$.

We embedded the blur estimation into a multiscale scheme.

Once we estimate $\hat{\mathbf{h}}$, we use a non-blind deconvolution algorithm to estimate the original sharp image.

5. Experimental Results

Test were run on a set of 4 test images with the 6 blur kernels.

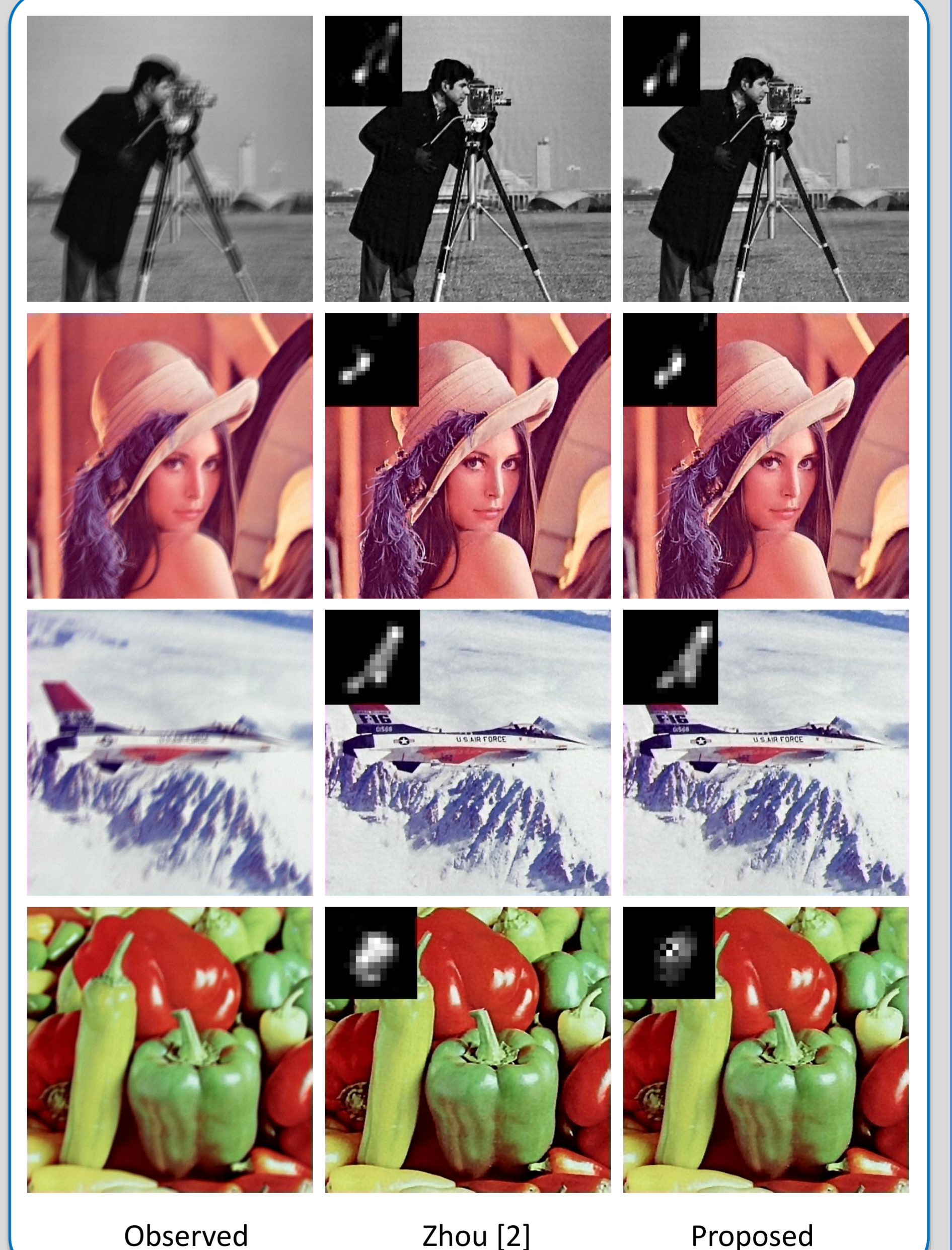


$\beta_\Gamma = 5000$, its real value, $\alpha_\Gamma, \pi_\Gamma$ were selected by a grid search.

We used 2 high-pass filters: $f_1 = [1, -1]$ and $f_2 = [1, -1]^T$ for the blur estimation. For the final image reconstruction, we also use the second order derivative filters $f_3 = [-1, 2, -1]$, $f_4 = [-1, 2, -1]^T$ and $f_5 = [1, -1; -1, 1]$. Comparison with the same settings was carried out with the method in [2].

PSNR values for the Y band

image	method	kernel					
		1	2	3	4	5	6
1	Proposed	31.04	30.56	31.18	33.37	31.35	30.60
	Zhou [9]	29.24	32.92	32.27	30.99	30.64	27.20
2	Proposed	31.03	30.99	31.41	32.63	30.81	32.04
	Zhou [9]	30.11	31.00	30.29	29.25	30.40	31.62
3	Proposed	29.55	31.07	30.64	31.28	29.20	24.47
	Zhou [9]	30.77	29.70	31.03	30.18	29.52	30.18
4	Proposed	31.04	30.47	31.75	31.66	30.31	31.21
	Zhou [9]	30.08	29.96	30.60	30.15	29.48	30.44



Observed

Zhou [2]

Proposed

Real images



Observed

Zhou [2]

Proposed