

1. Introduction

- Many engineering applications, such as attitude estimation, image processing, robotics, lead to models whose states are constrained to the Stiefel manifold $\mathcal{V}_{k,l}$.
- This work extends [1] in two ways:
- The observations are nonlinear functions of the state.
- We approximate the optimal importance function.

2. Problem Setup

• Let S_n denote the state of a system on the Stiefel manifold $\mathcal{V}_{k,l}$, i.e., $\{\mathbf{V} \in \mathbb{R}^{k \times l} : \mathbf{V}^T \mathbf{V} = \mathbf{I}_l\}, k > l$, according to

$$\mathbf{S}_{n}|\mathbf{S}_{n-1} \sim \mathbf{v}\mathbf{MF}(\mathbf{S}_{n}|\kappa\mathbf{S}_{n-1}) = \frac{\operatorname{etr}\left(\kappa\mathbf{S}_{n-1}^{T}\mathbf{S}_{n}\right)}{{}_{0}F_{1}\left(\frac{k}{2},\frac{\kappa^{2}}{4}\mathbf{I}_{l}\right)},$$

where $\kappa \in \mathbb{R}^+$ is a fixed hyperparameter and $_0F_1$ is the hypergeometric function with matrix argument.

• {S_n} gives rise to the observation sequence {Y_n}, Y_n $\in \mathbb{R}^{k \times l}$,

$$\mathbf{Y}_n | \mathbf{S}_n \sim \mathbf{N}_{k,l}(\mathbf{Y}_n | \mathbf{G}(\mathbf{S}_n), \mathbf{\Omega}, \mathbf{\Gamma}),$$

where $\mathbf{G} : \mathbb{R}^{k \times l} \to \mathbb{R}^{k \times l}$ is a possibly nonlinear function, and $\mathbf{N}_{k | l}$ is a matrix normal distribution on $\mathbb{R}^{k \times l}$.

• The particle filtering algorithm of [1] was restricted to $G(S_n) =$ S_n and used the prior importance function:

$$\mathbf{S}_{n}^{(q)} \sim p\left(\mathbf{S}_{n} | \mathbf{S}_{n-1}^{(q)}, \mathbf{Y}_{n}\right) = \mathbf{v} \mathbf{M} \mathbf{F}(\mathbf{S}_{n} | \kappa \mathbf{S}_{n-1}^{(q)}).$$

• The restriction on G can be trivially lifted, leading to the weight update equation

$$w_n^{(q)} \propto w_{n-1}^{(q)} \mathbf{N}_{k,l}(\mathbf{Y}_n | \mathbf{G}(\mathbf{S}_n^{(q)}), \mathbf{\Omega}, \mathbf{\Gamma}).$$

3. Proposed Method

Optimal importance function:

$$\mathbf{S}_{n}^{(q)} \sim p(\mathbf{S}_{n} | \mathbf{Y}_{n}, \mathbf{S}_{n-1}^{(q)}) = \frac{p(\mathbf{Y}_{n} | \mathbf{S}_{n}) p(\mathbf{S}_{n} | \mathbf{S}_{n-1}^{(q)})}{\int_{\mathcal{V}_{k,l}} p(\mathbf{Y}_{n} | \mathbf{S}_{n}) p(\mathbf{S}_{n} | \mathbf{S}_{n-1}^{(q)}) d\mathcal{V}_{k,l}(\mathbf{S}_{n})}.$$
(1)

• The integral in (1) cannot be analytically evaluated if $G(S_n)$ is a general nonlinear function. By linearizing $G(S_n)$ around S_{n-1} , we get

$$\mathbf{g}(\mathbf{s}_n) \approx \mathbf{g}(\mathbf{s}_{n-1}) + \mathbf{J}(\mathbf{s}_{n-1}) \left[\mathbf{s}_n - \mathbf{s}_{n-1}\right], \quad (2)$$

where $\mathbf{s}_n \triangleq \operatorname{vec}(\mathbf{S}_n)$, $\mathbf{g}(\mathbf{s}_n) \triangleq \operatorname{vec}(\mathbf{G}(\mathbf{S}_n))$ and $[\mathbf{J}(\mathbf{s}_{n-1})]_{kl} \triangleq$ $\left. \frac{\partial [\mathbf{g}(\mathbf{s})]_k}{\partial [\mathbf{s}]_l} \right|_{\mathbf{s}=\mathbf{s}_{n-1}}$ is a Jacobian matrix.

NONLINEAR STATE ESTIMATION USING PARTICLE FILTERS ON THE STIEFEL MANIFOLD

Claudio J. Bordin Jr., Universidade Federal do ABC, and Marcelo G. S. Bruno, Instituto Tecnológico de Aeronáutica.

• As a result of (2):

 $p(\mathbf{S}_n | \mathbf{Y}_n, \mathbf{S}_{n-1}^{(q)}) \approx \operatorname{FB}(\mathbf{S}_n | \mathbf{A}_n^{(q)}, \mathbf{B}_n^{(q)}),$

• The weights are then exactly propagated as

$$w_n^{(q)} \propto w_{n-1}^{(q)} \frac{\mathbf{v} \mathbf{MF}(\mathbf{S}_n^{(q)} | \kappa \mathbf{S}_{n-1}^{(q)}) \mathbf{N}_{k,l}(\mathbf{Y}_n | \mathbf{G}(\mathbf{S}_n^{(q)}), \mathbf{\Omega}, \mathbf{\Gamma})}{\mathbf{FB}(\mathbf{S}_n^{(q)} | \mathbf{A}_n^{(q)}, \mathbf{B}_n^{(q)})}.$$
 (3)

where FB stands for the matrix Fisher-Bingham p.d.f. on $\mathcal{V}_{n,k}$:

$$FB(\mathbf{S}_n | \mathbf{A}_n, \mathbf{B}_n) = \frac{\exp\left\{ \operatorname{tr}(\mathbf{A}_n^T \mathbf{S}_n) + \operatorname{vec}(\mathbf{S}_n)^T \mathbf{B}_n \operatorname{vec}(\mathbf{S}_n) \right\}}{c_{FB}(\mathbf{A}_n, \mathbf{B}_n)},$$

where $c_{FB}(\mathbf{A}_n, \mathbf{B}_n)$ is the matrix Fisher-Bingham p.d.f. normalization constant, and

$$\mathbf{A}_{n} \triangleq \operatorname{vec}^{-1}(\tilde{\mathbf{y}}_{n}^{T} \boldsymbol{\Sigma}^{-1} \mathbf{J}(\mathbf{s}_{n-1})) + \kappa \mathbf{S}_{n-1},$$

$$\mathbf{B}_{n} \triangleq -\frac{1}{2} \mathbf{J}(\mathbf{s}_{n-1})^{T} \boldsymbol{\Sigma}^{-1} \mathbf{J}(\mathbf{s}_{n-1}),$$

$$\tilde{\mathbf{y}}_{n} \triangleq \mathbf{y}_{n} - \mathbf{g}(\mathbf{s}_{n-1}) + \mathbf{J}(\mathbf{s}_{n-1})\mathbf{s}_{n-1},$$

$$\boldsymbol{\Sigma} \triangleq \boldsymbol{\Gamma} \otimes \boldsymbol{\Omega}.$$

4. Sampling from a matrix Fisher-Bingham p.d.f.

- We adapted from the algorithm in [2, Sec. 3.3], originally developed for the matrix Bingham-Von Mises-Fisher p.d.f.
- Under the restriction that \mathbf{B}_n is a block-diagonal matrix,

$$FB(\mathbf{S}_n|\mathbf{A}_n, \mathbf{B}_n) \propto \prod_{m=1}^{l} \exp\left(\mathbf{A}_n[, m]^T \mathbf{S}_n[, m] + \mathbf{S}_n[, m]^T \mathbf{B}_n(m) \mathbf{S}_n[, m]\right),$$

where [,m] stands for the m-th column of a matrix and $\mathbf{B}(m) \in \mathbb{R}$ $\mathbb{R}^{k \times k}$ denotes the *m*-th block of the diagonal of \mathbf{B}_n .

- As the columns of S_n are orthogonal with probability 1, $S_n = 1$ $[Nz \ S_n[,-1]]$, where $S_n[,-1]$ is the matrix formed by removing the first column of \mathbf{S}_n , $\mathbf{N} \in \mathbb{R}^{k \times (k-l+1)}$ is an orthonormal basis for the null space of $S_n[,-1]$, and z is a (k - l + 1) unit-norm column vector.
- The conditional p.d.f. of z is then given by [2]

$$p(\mathbf{z}|\mathbf{S}_n[,-1]) \propto \exp\left(\mathbf{A}_n[,1]^T\mathbf{N}\mathbf{z} + \mathbf{z}^T\mathbf{N}^T\mathbf{B}_n(1)\mathbf{N}\mathbf{z}\right)$$

which is a Fisher-Bingham density on the unit sphere.

- A Markov chain with stationary p.d.f. $FB(\mathbf{S}_n | \mathbf{A}_n, \mathbf{B}_n)$ can be obtained via the Gibbs sampler:
- -Given $S_n^{<j>} = S$, the *j*-th element of the chain, compute steps 1 to 4 for each $m \in \{1, \ldots, l\}$ in a random order:
- 1. compute N, an orthonormal basis for the null space of S[, -m];
- 2. compute $\tilde{\mathbf{a}} = \mathbf{A}_n[,m]^T \mathbf{N}$ and $\tilde{\mathbf{B}} = \mathbf{N}^T \mathbf{B}_n(m) \mathbf{N};$
- 3. sample z from a Fisher-Bingham density on the unit sphere with parameters $\tilde{\mathbf{a}}$ and $\tilde{\mathbf{B}}$. $1 \cot \mathbf{S}[m] - \mathbf{N}\mathbf{z}$

4. Set
$$\mathbf{S}[,m] = \mathbf{N}\mathbf{Z}$$

-Set $\mathbf{S}_n^{<j+1>} = \mathbf{S}$.

 $c_{FB}($

 $c_{MF}(k\mathbf{S}_{n-1}) = c_{FB}(k\mathbf{S}_{n-1}, \mathbf{0}).$

1. Ge
$$\mathcal{V}_k$$

where $S_n^{<i>}$ denotes the *i*-th estimate of the weighted average, with $S_n^{<0>}$ chosen as a random element of the particle set, and \mathcal{M} and \mathcal{M}^{-1} are the orthographic retraction and lifting maps [3].

particles.

5. Computation of the matrix Fisher-Bingham p.d.f. normalization constant

• To update the weights (Equation 3), it is necessary to compute the normalization constants

$$(\mathbf{A}_n, \mathbf{B}_n) \triangleq \int_{\mathcal{V}_{k,l}} \exp\left\{ \operatorname{tr}(\mathbf{A}_n^T \mathbf{S}) + \operatorname{vec}(\mathbf{S})^T \mathbf{B}_n \operatorname{vec}(\mathbf{S}) \right\} d\mathcal{V}_{k,l}(\mathbf{S}),$$
(4)

existing approaches did not perform adequately, we oduced the Quasi Monte Carlo algorithm:

enerate low-discrepancy samples uniformly distributed on

2. Approximate (4) as

$$\operatorname{Vol}(\mathcal{V}_{k,l}) \frac{1}{N_S} \sum_{i=1}^{N_S} \left\{ \operatorname{tr}(\mathbf{A}_n^T \mathbf{X}^{}) + \operatorname{vec}(\mathbf{X}^{})^T \mathbf{B}_n \operatorname{vec}(\mathbf{X}^{}) \right\}$$

where N_S is the number of samples, $\mathbf{X}^{< i>}$ is the *i*-th generated sample, and $Vol(\mathcal{V}_{k,l})$ is the volume of $\mathcal{V}_{k,l}$.

6. Computation of the weighted averages on the Stiefel manifold

• Ideally, one would estimate the state as a Karcher mean, i.e., the value of $\hat{\mathbf{S}}_n$ that minimizes the weighted mean square geodesic distance to the particle set.

• To reduce computational complexity, we evaluated the weighted averages over the Stiefel manifold as

$$\mathbf{S}_{n}^{} = \mathcal{M}_{\mathbf{S}_{n}^{}} \left(\sum_{q=1}^{Q} w_{n}^{(q)} \ \mathcal{M}_{\mathbf{S}_{n}^{}}^{-1} \left(\mathbf{S}_{n}^{(q)} \right) \right), \ i \ge 0,$$

7. Numerical Experiment

• We performed numerical simulations with 150 independent trials of 100 synthetic data samples. Particle filters used 300

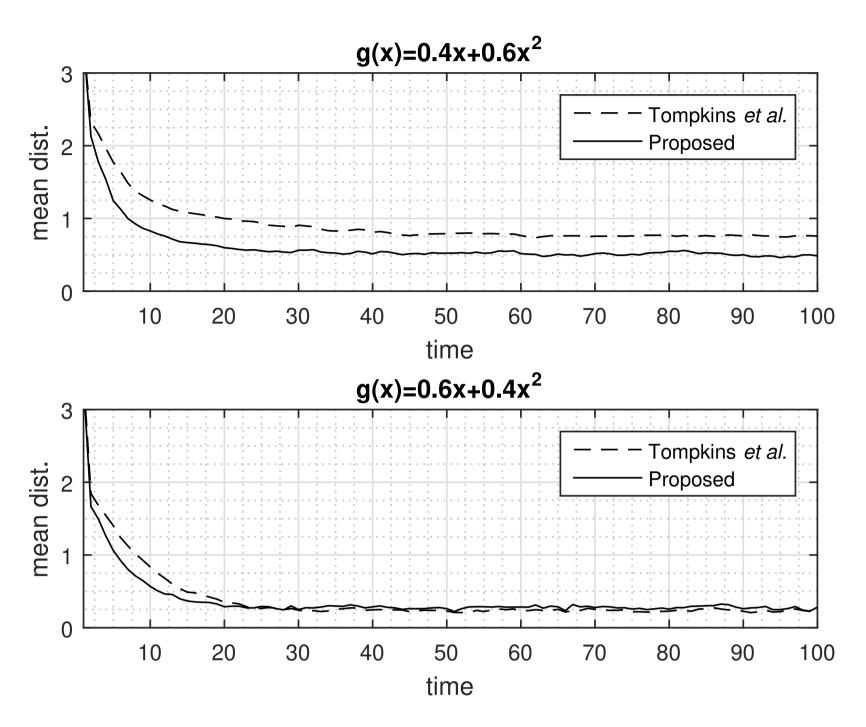


Figure 1: Mean geodesic distance for the proposed algorithm and that of [1] as a function of time, for distinct nonlinear observation functions q(x).

complexity.

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• We assumed that G acts element-wise, so that the Jacobian $J(s_{n-1})$ is diagonal. The parameters were set to $\kappa = 150$, $\Omega \triangleq \mathbf{I}_l, \Gamma \triangleq \mathbf{I}_k \sigma^2$, and $\Sigma = \mathbf{I}_{kl} \sigma^2$, with $\sigma^2 = 0.05, k = 3, l = 2$. • The algorithms' performance was evaluated in terms of the mean geodesic distance from the true state S_n to the estimated state $\hat{\mathbf{S}}_n$, i.e., $d(\mathbf{S}_n, \hat{\mathbf{S}}_n) = \|\operatorname{Exp}_{\mathbf{S}_n}^{-1}(\hat{\mathbf{S}}_n)\|_F$.

• For stronger nonlinearity (top), the proposed method exhibited an asymptotic error about 30% smaller than the method of [1]

8. Conclusions

• For certain choices of G, the proposed method outperforms that of [1] at the expense of increased computational

• Most of the computational complexity of the proposed method is related to drawing samples from and computing normalization constants for the matrix Fisher-Bingham density.

References

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