

Choosing the diagonal loading factor for linear signal estimation using cross validation

**Jun Tong, Qinghua Guo, Jiangtao Xi, Yanguang Yu¹
Peter J. Schreier²**

¹School of Electrical, Computer & Telecommunication Engineering
The University of Wollongong, Australia
jtong@uow.edu.au

²Signal and System Theory Group
Universität Paderborn, Germany
peter.schreier@sst.upb.de

ICASSP 2016

- Consider a system with signal (input) \mathbf{x} and measurement \mathbf{y}
- Linear minimum mean-squared error (LMMSE) estimator

$$\hat{\mathbf{x}} = \mathbf{C}_{yx}^\dagger \mathbf{C}_{yy}^{-1} \mathbf{y}$$

minimizes

$$\text{MSE}_x \triangleq \mathbb{E}_x[\|\mathbf{x} - \hat{\mathbf{x}}\|^2]$$

- In practice, **sample covariance matrices** (SCMs) computed from length- T training data:

$$\hat{\mathbf{C}}_{yy} = \frac{1}{T} \mathbf{Y}\mathbf{Y}^\dagger, \quad \hat{\mathbf{C}}_{yx} = \frac{1}{T} \mathbf{Y}\mathbf{X}^\dagger$$

- With **low sample support**, LMMSE estimator may perform poorly due to **model mismatch**

- Consider a system with signal (input) \mathbf{x} and measurement \mathbf{y}
- Linear minimum mean-squared error (LMMSE) estimator

$$\hat{\mathbf{x}} = \mathbf{C}_{yx}^\dagger \mathbf{C}_{yy}^{-1} \mathbf{y}$$

minimizes

$$\text{MSE}_{\mathbf{x}} \triangleq \mathbb{E}_{\mathbf{x}}[\|\mathbf{x} - \hat{\mathbf{x}}\|^2]$$

- In practice, **sample covariance matrices** (SCMs) computed from length- T training data:

$$\hat{\mathbf{C}}_{yy} = \frac{1}{T} \mathbf{Y} \mathbf{Y}^\dagger, \quad \hat{\mathbf{C}}_{yx} = \frac{1}{T} \mathbf{Y} \mathbf{X}^\dagger$$

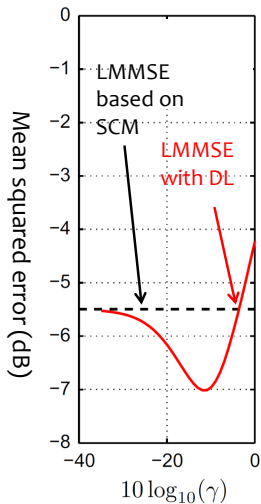
- With **low sample support**, LMMSE estimator may perform poorly due to **model mismatch**

- Robustness can be improved by diagonal loading (DL) with **diagonal loading factor** (DLF) $\gamma \geq 0$:

$$\hat{\mathbf{x}} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

- A.k.a. **Tikhonov regularization** or **ridge regression**
- Improves **condition number** of the matrix to be inverted
- Achieves a better **bias-variance trade-off** \Rightarrow lower MSE

How do we choose the DLF?



- DLF γ significantly affects performance
- Typical ad-hoc choice:

$$\gamma = 10\lambda_{\min}$$

- Need methods to automatically tune the DLF



Objective of this work

Given the estimated covariance matrices, automatically choose the optimal γ for

$$\hat{\mathbf{x}}_{\gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

such that the MSE of estimating \mathbf{x} is minimized:

$$\gamma^* = \arg \min_{\gamma} E_{\mathbf{x}} [\|\mathbf{x} - \hat{\mathbf{x}}_{\gamma}\|^2]$$

- A more general problem: Optimize the shrinkage factors (α, γ) for the estimate

$$\hat{\mathbf{x}}_{\alpha, \gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\alpha \hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

which reduces to DL for $\alpha = 1$

Objective of this work

Given the estimated covariance matrices, automatically choose the optimal γ for

$$\hat{\mathbf{x}}_{\gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

such that the MSE of estimating \mathbf{x} is minimized:

$$\gamma^* = \arg \min_{\gamma} E_{\mathbf{x}} [\|\mathbf{x} - \hat{\mathbf{x}}_{\gamma}\|^2]$$

- A more general problem: Optimize the shrinkage factors (α, γ) for the estimate

$$\hat{\mathbf{x}}_{\alpha, \gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\alpha \hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

which reduces to DL for $\alpha = 1$

Techniques based on **random matrix theory** (RMT) and **large system assumption**

① Optimize **estimation of the covariance matrix**

- **Examples:** Ledoit and Wolf (J. Multivariate Analysis 2004), Stoica et al. (TSP 2008), Chen et al. (TSP 2010)
- Achieves near-optimal covariance matrix estimation
- But generally suboptimal for **signal estimation**

② Maximize **SINR**

- **Examples:** Mestre and Lagunas (TSP 2006), Zhang et al. (TSP 2013)
- Generally suboptimal for minimizing MSE

③ Minimize **MSE**

- **Examples:** Wen et al. (SPL 2013), Zhang et al. (TSP 2013)
- based on SCM
- do not account for **differently distributed** training and application data

Techniques based on **random matrix theory** (RMT) and **large system assumption**

1 Optimize **estimation of the covariance matrix**

- **Examples:** Ledoit and Wolf (J. Multivariate Analysis 2004), Stoica et al. (TSP 2008), Chen et al. (TSP 2010)
- Achieves near-optimal covariance matrix estimation
- But generally suboptimal for **signal estimation**

2 Maximize **SINR**

- **Examples:** Mestre and Lagunas (TSP 2006), Zhang et al. (TSP 2013)
- Generally suboptimal for minimizing MSE

3 Minimize **MSE**

- **Examples:** Wen et al. (SPL 2013), Zhang et al. (TSP 2013)
- based on SCM
- do not account for **differently distributed** training and application data

Techniques based on **random matrix theory** (RMT) and **large system assumption**

① Optimize **estimation of the covariance matrix**

- **Examples:** Ledoit and Wolf (J. Multivariate Analysis 2004), Stoica et al. (TSP 2008), Chen et al. (TSP 2010)
- Achieves near-optimal covariance matrix estimation
- But generally suboptimal for **signal estimation**

② Maximize **SINR**

- **Examples:** Mestre and Lagunas (TSP 2006), Zhang et al. (TSP 2013)
- Generally suboptimal for minimizing MSE

③ Minimize **MSE**

- **Examples:** Wen et al. (SPL 2013), Zhang et al. (TSP 2013)
- based on SCM
- do not account for **differently distributed** training and application data

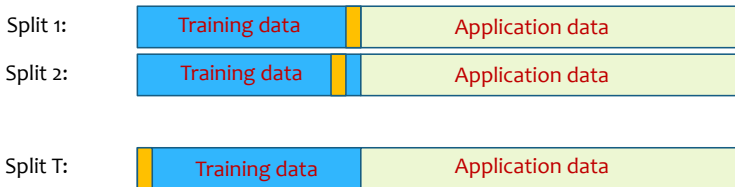
- We propose choosing the DLF based on **cross-validation** (CV)
- We derive **computationally efficient** calculation schemes
- Not based on random matrix theory
- Explicitly target the **minimization of the MSE** for signal estimation
- Allow **different distributions** for training and application data

Leave-one-out cross validation (LOOCV)

- Choosing γ is a **model selection problem**
- Assume first that the training and application data are identically distributed
- Reserve some of the training data for model validation **under the signal estimation criterion**:



- LOOCV splits repeatedly, reserving **one symbol** for validation each time:



Direct implementation of LOOCV is expensive

Using SCMs and T samples, LOOCV chooses

$$\gamma^* = \arg \min_{\gamma} \frac{1}{T} \sum_{i=1}^T \|\mathbf{x}_i - \mathbf{W}_{\sim i, \gamma}^{\dagger} \mathbf{y}_i\|^2$$

with $\mathbf{W}_{\sim i, \gamma} = \left(\mathbf{Y}_{\sim i} \mathbf{Y}_{\sim i}^{\dagger} + \gamma \mathbf{I} \right)^{-1} \mathbf{Y}_{\sim i} \mathbf{X}_{\sim i}^{\dagger}$

- If we test K candidates for γ , this requires KT matrix inversions
- For N -dimensional \mathbf{y} , the resulting **complexity** $O(KTN^3)$ can be prohibitive

Direct implementation of LOOCV is expensive

Using SCMs and T samples, LOOCV chooses

$$\gamma^* = \arg \min_{\gamma} \frac{1}{T} \sum_{i=1}^T \|\mathbf{x}_i - \mathbf{W}_{\sim i, \gamma}^{\dagger} \mathbf{y}_i\|^2$$

with $\mathbf{W}_{\sim i, \gamma} = \left(\mathbf{Y}_{\sim i} \mathbf{Y}_{\sim i}^{\dagger} + \gamma \mathbf{I} \right)^{-1} \mathbf{Y}_{\sim i} \mathbf{X}_{\sim i}^{\dagger}$

- If we test K candidates for γ , this requires KT matrix inversions
- For N -dimensional \mathbf{y} , the resulting **complexity** $O(KTN^3)$ can be prohibitive

- For SCMs, we apply the **Woodbury matrix identity** to simplify the problem to

$$\gamma^* = \arg \min_{\gamma} \left\| \mathbf{X} - \mathbf{X} (\mathbf{B}_{\gamma} - \mathbf{D}_{\mathbf{B}_{\gamma}}) (\mathbf{I} - \mathbf{D}_{\mathbf{B}_{\gamma}})^{-1} \right\|^2$$

where

$$\mathbf{B}_{\gamma} \triangleq \mathbf{Y}^{\dagger} (\mathbf{Y}\mathbf{Y}^{\dagger} + \gamma\mathbf{I})^{-1} \mathbf{Y}$$

and $\mathbf{D}_{\mathbf{B}_{\gamma}}$ is a diagonal matrix with diagonal entries of \mathbf{B}_{γ}

- This is a **univariate optimization problem**, which can be solved using standard tools
- Computing the **SVD** of \mathbf{Y} can further **accelerate** the evaluation of the cost function for different candidates γ

- For SCMs, we apply the **Woodbury matrix identity** to simplify the problem to

$$\gamma^* = \arg \min_{\gamma} \left\| \mathbf{X} - \mathbf{X} (\mathbf{B}_{\gamma} - \mathbf{D}_{\mathbf{B}_{\gamma}}) (\mathbf{I} - \mathbf{D}_{\mathbf{B}_{\gamma}})^{-1} \right\|^2$$

where

$$\mathbf{B}_{\gamma} \triangleq \mathbf{Y}^{\dagger} (\mathbf{Y}\mathbf{Y}^{\dagger} + \gamma\mathbf{I})^{-1} \mathbf{Y}$$

and $\mathbf{D}_{\mathbf{B}_{\gamma}}$ is a diagonal matrix with diagonal entries of \mathbf{B}_{γ}

- This is a **univariate optimization problem**, which can be solved using standard tools
- Computing the **SVD** of \mathbf{Y} can further **accelerate** the evaluation of the cost function for different candidates γ

- **Training** and **application data** may have **different distributions** (e.g., orthogonal training)
- In this case, we exploit **spatial correlation** between entries of \mathbf{y}
 - E.g., in the MIMO channel model $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{z}$, this correlation is introduced by \mathbf{H}
- **Assumption:** Estimates of covariance matrices ($\hat{\mathbf{C}}_{yx}$, $\hat{\mathbf{C}}_{yy}$) available
- Perform **spatial LOOCV** on the **application data**. That is, choose γ to **minimize the MSE** of predicting $y_d^{(n)}$ from $\mathbf{y}_d^{(\sim n)}$:

$$\gamma^* = \arg \min_{\gamma} \frac{1}{ND} \sum_{d=1}^D \sum_{n=1}^N \left| y_d^{(n)} - \hat{y}_{d,\gamma}^{(n)} \right|^2$$

with length- D application data of dimension N

- **Training** and **application data** may have **different distributions** (e.g., orthogonal training)
- In this case, we exploit **spatial correlation** between entries of **y**
 - E.g., in the MIMO channel model $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{z}$, this correlation is introduced by **H**
- **Assumption:** Estimates of covariance matrices ($\hat{\mathbf{C}}_{yx}$, $\hat{\mathbf{C}}_{yy}$) available
- Perform **spatial LOOCV** on the **application data**. That is, choose γ to **minimize the MSE** of predicting $y_d^{(n)}$ from $\mathbf{y}_d^{(\sim n)}$:

$$\gamma^* = \arg \min_{\gamma} \frac{1}{ND} \sum_{d=1}^D \sum_{n=1}^N \left| y_d^{(n)} - \hat{y}_{d,\gamma}^{(n)} \right|^2$$

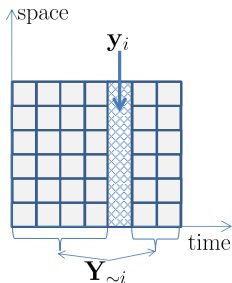
with length- D application data of dimension N

- **Training** and **application data** may have **different distributions** (e.g., orthogonal training)
- In this case, we exploit **spatial correlation** between entries of \mathbf{y}
 - E.g., in the MIMO channel model $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{z}$, this correlation is introduced by \mathbf{H}
- **Assumption:** Estimates of covariance matrices ($\hat{\mathbf{C}}_{yx}$, $\hat{\mathbf{C}}_{yy}$) available
- Perform **spatial LOOCV** on the **application data**. That is, choose γ to **minimize the MSE** of predicting $y_d^{(n)}$ from $\mathbf{y}_d^{(\sim n)}$:

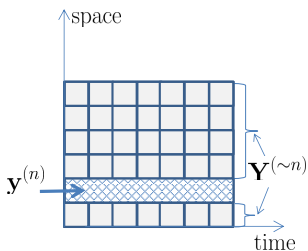
$$\gamma^* = \arg \min_{\gamma} \frac{1}{ND} \sum_{d=1}^D \sum_{n=1}^N \left| y_d^{(n)} - \hat{y}_{d,\gamma}^{(n)} \right|^2$$

with length- D application data of dimension N

- Perform **spatial LOOCV** on the **application data**:



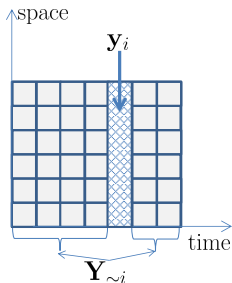
Split the **training**
data w.r.t. **time**



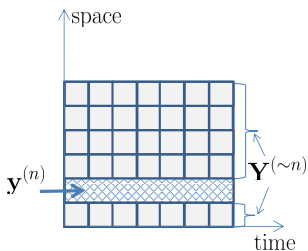
Split the **application**
data w.r.t. **space**

- As before, **computationally efficient** implementations are derived using the Woodbury matrix identity

- Perform **spatial LOOCV** on the **application data**:



Split the **training**
data w.r.t. **time**



Split the **application**
data w.r.t. **space**

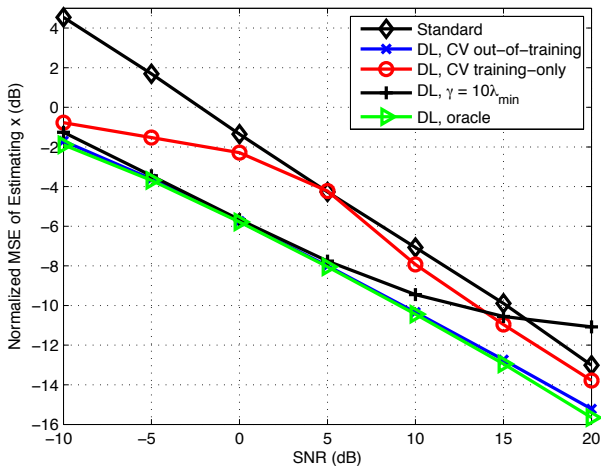
- As before, **computationally efficient** implementations are derived using the Woodbury matrix identity

Results: MSE vs. SNR with orthogonal training

20×20 MIMO channel model: $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{z}$

Training: orthogonal (DFT), $T = 24$

Application data: Gaussian, $D = 24$

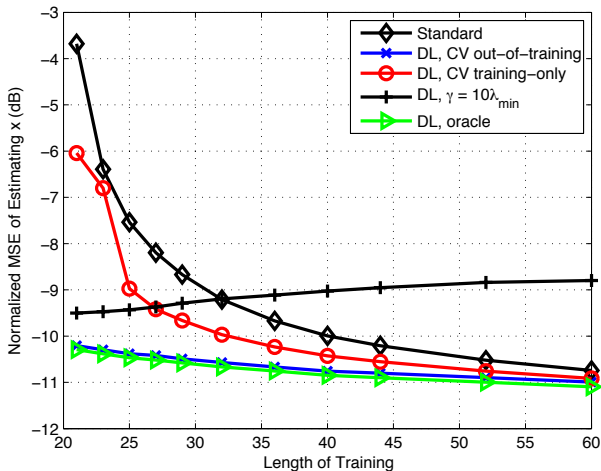


Results: MSE vs. T with orthogonal training

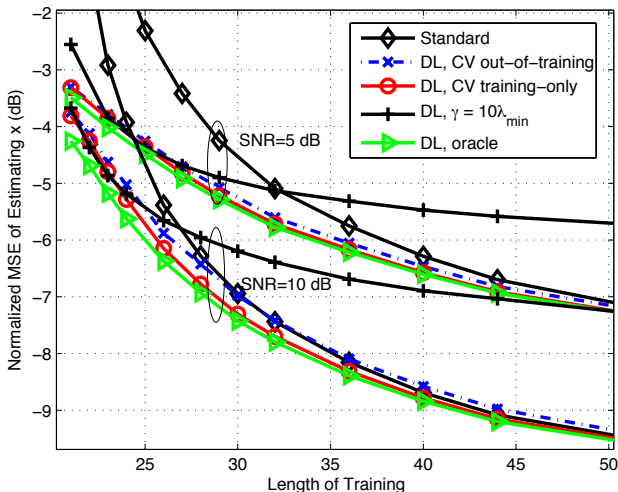
Training: orthogonal (DFT), varying T

Application data: Gaussian, $D = 24$

SNR: 10 dB



Application and training data identically distributed



- LOOCV can be **efficiently** used to choose the **DLF minimizing MSE**
- Can handle both **identically distributed** and **differently distributed** training and application data, by splitting w.r.t. **time** and **space**, respectively
- These ideas can be generalized to the shrinkage estimator

$$\hat{\mathbf{x}}_{\alpha, \gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\alpha \hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

- See our forthcoming journal paper:
J. Tong, P. J. Schreier, Q. Guo, S. Tong, J. Xi, and Y. Yu,
“Shrinkage of covariance matrices for linear signal estimation using cross-validation,” to appear in *IEEE Trans. Signal Processing*

- LOOCV can be **efficiently** used to choose the **DLF minimizing MSE**
- Can handle both **identically distributed** and **differently distributed** training and application data, by splitting w.r.t. **time** and **space**, respectively
- These ideas can be generalized to the shrinkage estimator

$$\hat{\mathbf{x}}_{\alpha, \gamma} = \hat{\mathbf{C}}_{yx}^{\dagger} \left(\alpha \hat{\mathbf{C}}_{yy} + \gamma \mathbf{I} \right)^{-1} \mathbf{y}$$

- See our forthcoming journal paper:
J. Tong, P. J. Schreier, Q. Guo, S. Tong, J. Xi, and Y. Yu,
“Shrinkage of covariance matrices for linear signal estimation using cross-validation,” to appear in *IEEE Trans. Signal Processing*