

Fast variational Bayesian signal recovery in the presence of Poisson-Gaussian Noise.

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Mixed Poisson Gaussian noise

General model

We observe $\mathbf{y} \in \mathbb{R}^N$ according to the following model

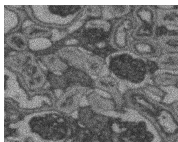
$$\mathbf{y} = \mathbf{z} + \mathbf{w}.$$

- $\mathbf{z} \in \mathbb{R}^N$ \rightsquigarrow Poisson noise ($\mathbf{z} \sim \mathcal{P}(\mathbf{H}\mathbf{x})$)
- $\mathbf{x} \in \mathbb{R}^Q$ \rightsquigarrow unknown original signal
- $\mathbf{H} \in \mathbb{R}^{N \times Q}$ \rightsquigarrow observation operator
- $\mathbf{w} \in \mathbb{R}^N$ \rightsquigarrow additive Gaussian noise ($\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \sigma^2)$)

WHERE?

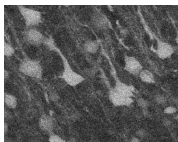
- ▶ CCD camera images [Healey et al. 1994]
- ▶ Medical images [Nichols et al. 2002]
- ▶ Biological images (fluorescence microscopy) [Pawley 1994]
- ▶ Astronomical images [Benvenuto et al. 2008]

Mixed Poisson Gaussian noise



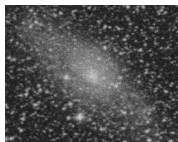
$$x_{\max} = 20$$
$$\sigma^2 = 9$$

H blur with PSF
 h : Uniform 5×5



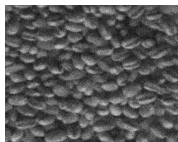
$$x_{\max} = 60$$
$$\sigma^2 = 36$$

H blur with PSF
 h : Gaussian 9×9
std 0.5



$$x_{\max} = 100$$
$$\sigma^2 = 36$$

H blur with PSF
 h : Uniform 3×3



$$x_{\max} = 150$$
$$\sigma^2 = 40$$

H blur with PSF
 h : Gaussian 7×7
std 1

OBJECTIVE?

↪ Provide an estimate $\hat{\mathbf{x}}$ of \mathbf{x} from the collected data \mathbf{y} .

Bayesian formulation

OBSERVATION MODEL: $\mathbf{y} = \mathbf{z} + \mathbf{w}$, $\mathbf{z} \sim \mathcal{P}(\mathbf{H}\mathbf{x})$, $\mathbf{w} \sim \mathcal{N}(0, \sigma^2)$

$$p(\mathbf{y} | \mathbf{x}) = \prod_{i=1}^M \left(\sum_{n=1}^{+\infty} \frac{e^{-[\mathbf{H}\mathbf{x}]_i} ([\mathbf{H}\mathbf{x}]_i)^n}{n!} \frac{e^{-\frac{1}{2\sigma^2}(y_i-n)^2}}{\sqrt{2\pi\sigma^2}} \right)$$

PRIOR DISTRIBUTIONS:

$$p(\mathbf{x} | \gamma) = \tau \gamma^{\frac{N}{2\kappa}} \exp \left(-\gamma \sum_{j=1}^Q \|\mathbf{D}_j \mathbf{x}\|^{2\kappa} \right) \quad p(\gamma) \propto \gamma^{\alpha-1} \exp(-\beta\gamma)$$

BAYES FORMULA: $p(\mathbf{x} | \mathbf{y}, \gamma) \propto p(\mathbf{x} | \gamma)p(\mathbf{y} | \mathbf{x})p(\gamma)$.

✗ DIFFICULTIES: $p(\mathbf{x} | \mathbf{y}, \gamma)$ of a complicated form

- ▶ MAP estimation: tuning parameters
- ▶ Posterior mean: intractable

✓ SOLUTIONS:

- ▶ MCMC methods: computationally expensive
- ▶ Variational Bayesian (VB) approaches

Variational Bayesian methodology

Let us denote Θ the set of the unknown parameters

PRINCIPLE: provide a separable approximation

$q(\Theta) = \prod_{j=1}^J q(\Theta_j)$ of the true posterior distribution $p(\Theta | \mathbf{y})$

OPTIMIZATION PROBLEM:

$$q^{opt} = \underset{q}{\operatorname{argmin}} \mathcal{KL}(q(\Theta) \| p(\Theta | \mathbf{y})) \quad \text{s.t.} \quad q \text{ is a p.d.f}$$

where

$$\mathcal{KL}(q(\Theta) \| p(\Theta | \mathbf{y})) = \int q(\Theta) \frac{q(\Theta)}{p(\Theta | \mathbf{y})} d\Theta$$

An analytical solution (classical variational Bayesian methods) is :

$$(\forall j \in \{1, \dots, J\}) \quad q(\Theta_j) \propto \exp \left(\langle p(\mathbf{y}, \Theta) \rangle_{\prod_{i \neq j} q(\Theta_i)} \right),$$

where $\langle \cdot \rangle_{\prod_{i \neq j} q(\Theta_i)} = \int \cdot \prod_{i \neq j} q(\Theta_i) d\Theta_i$

Assumptions and approximations

- **SEPARABILITY**: We assume that

$$q(\Theta) = q(\mathbf{x})q(\gamma)$$

- **GENERALIZED ANSCOMBE TRANSFORM (GAST) APPROXIMATION**

The likelihood of vector $\tilde{\mathbf{y}} \in \mathbb{R}^M$ with components $(\tilde{y}_i)_{1 \leq i \leq M} = 2(\sqrt{y_i + \delta})_{1 \leq i \leq M}$, where $\delta = \frac{3}{8} + \sigma^2$, is approximately given by

$$p(\tilde{\mathbf{y}} | \mathbf{x}) = \prod_{i=1}^M \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2} \left(\tilde{y}_i - 2\sqrt{[\mathbf{H}\mathbf{x}]_i + \delta}\right)^2\right)$$

- ✗ **DIFFICULTIES**: $p(\mathbf{x}, \gamma | \mathbf{y})$ is of a complicated form
- ✓ **SOLUTIONS**: Majorizing approximations

Construction of the majorizing approximations

- MAJORANT FUNCTION OF THE NEG-LOG LIKELIHOOD :

Let $\mathbf{w} = (w_i)_{1 \leq i \leq M} \in [0, +\infty)^M$.

Since the function $t \mapsto \sqrt{t + \delta}$ is **concave** with a **Lipschitz continuous gradient**, then $T(\tilde{\mathbf{y}}, \mathbf{x}; \mathbf{w}) = \sum_{i=1}^M T_i(\tilde{y}_i, [\mathbf{H}\mathbf{x}]_i; w_i)$

where, for every $i \in \{1, \dots, M\}$,

$$\begin{aligned} T_i(\tilde{y}_i, [\mathbf{H}\mathbf{x}]_i; w_i) &= \frac{1}{2} \tilde{y}_i^2 + 2([\mathbf{H}\mathbf{x}]_i + \delta) - 2\tilde{y}_i \sqrt{w_i + \delta} \\ &\quad - \tilde{y}_i (w_i + \delta)^{-\frac{1}{2}} ([\mathbf{H}\mathbf{x}]_i - w_i) \\ &\quad + \frac{1}{4} \delta^{-\frac{3}{2}} \tilde{y}_i ([\mathbf{H}\mathbf{x}]_i - w_i)^2. \end{aligned}$$

is a majorant of $-\log p(\tilde{\mathbf{y}} | \mathbf{x})$.

Construction of the majorizing approximations

- MAJORANT FUNCTION OF THE NEG-LOG PRIOR OF \mathbf{x} :

Let $\boldsymbol{\lambda} = (\lambda_j)_{1 \leq j \leq Q} \in [0, +\infty)^Q$.

Using the following inequality:

$$(\forall v \geq 0)(\forall \nu > 0) \quad v^\kappa \leq (1 - \kappa)v^\kappa + \kappa v^{\kappa-1}\nu,$$

then $Q(\mathbf{x}, \gamma; \boldsymbol{\lambda}) = \sum_{j=1}^Q Q_j(\mathbf{D}_j \mathbf{x}, \gamma; \lambda_j)$ where for every $j \in \{1, \dots, Q\}$,

$$(\forall j \in \{1, \dots, Q\}) \quad Q_j(\mathbf{D}_j \mathbf{x}, \gamma; \lambda_j) = \gamma \frac{\kappa \|\mathbf{D}_j \mathbf{x}\|^2 + (1 - \kappa)\lambda_j}{\lambda_j^{1-\kappa}}.$$

is a majorant of $-\log p(\mathbf{x})$.

Construction of the majorizing approximations

Let

$$L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda) = C(\tilde{\mathbf{y}}) \exp[-T(\tilde{\mathbf{y}}, \mathbf{x}; \mathbf{w}) - Q(\mathbf{x}, \gamma; \lambda)] p(\gamma)$$

$$\text{where } C(\tilde{\mathbf{y}}) = p(\tilde{\mathbf{y}})^{-1} (2\pi)^{-M/2} \tau \gamma^{\frac{N}{2\kappa}}$$

Then

- ▶ $p(\Theta | \tilde{\mathbf{y}}) \geq L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda)$
- ▶ $\mathcal{KL}(q(\Theta) || p(\Theta | \tilde{\mathbf{y}})) \leq \mathcal{KL}(q(\Theta) || L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$

To recap...

Let $\Theta(\mathbf{x}, \gamma)$, $q(\Theta) = q(\mathbf{x})q(\gamma)$

OPTIMIZATION PROBLEM:

$$q(\Theta) = \underset{q}{\operatorname{argmin}} \mathcal{KL}(q(\Theta) \| p(\Theta | \tilde{\mathbf{y}})) \quad (1)$$

MAJORATION: $\mathcal{KL}(q(\Theta) \| p(\Theta | \tilde{\mathbf{y}})) \leq \mathcal{KL}(q(\Theta) \| L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$

Then Thus, Problem (1) can be solved by alternating the following steps:

- ▶ Mimimize $\mathcal{KL}(q(\Theta) \| L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$ w.r.t. $q(\mathbf{x})$,
- ▶ Update the auxiliary variable \mathbf{w} in order to minimize $\mathcal{KL}(q(\Theta) \| L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$,
- ▶ Update the auxiliary variable λ in order to minimize $\mathcal{KL}(q(\Theta) \| L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$,
- ▶ Mimimize the $\mathcal{KL}(q(\Theta) \| L(\Theta | \tilde{\mathbf{y}}; \mathbf{w}, \lambda))$ w.r.t. $q(\gamma)$.

Updating $q(\mathbf{x})$

$$\begin{aligned}
 q^{k+1}(\mathbf{x}) &\propto \exp \left(\left\langle \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^k, \boldsymbol{\lambda}^k) \right\rangle_{q^k(\gamma)} \right) \\
 &\propto \exp \left(\int \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^k, \boldsymbol{\lambda}^k) q^k(\gamma) d\gamma \right) \\
 &= \mathcal{N}(\mathbf{x}; \mathbf{m}_{k+1}, \boldsymbol{\Sigma}_{k+1})
 \end{aligned}$$

$$\begin{cases}
 \boldsymbol{\Sigma}_{k+1}^{-1} &= \frac{1}{2} \delta^{-\frac{3}{2}} \mathbf{H}^\top \text{Diag}(\tilde{\mathbf{y}}) \mathbf{H} + 2 \mathbb{E}_{q^k(\gamma)} \mathbf{D}^\top \boldsymbol{\Lambda}^k \mathbf{D}, \\
 \mathbf{m}_{k+1} &= \boldsymbol{\Sigma}_{k+1} \mathbf{H}^\top \mathbf{u},
 \end{cases} \quad (1)$$

where

- $\mathbf{u} = (u_i)_{1 \leq i \leq M}$, $u_i = \tilde{y}_i (w_i^k + \delta)^{-\frac{1}{2}} + \frac{1}{2} \tilde{y}_i \delta^{-\frac{3}{2}} w_i^k - 2$
- $\boldsymbol{\Lambda}$ is the diagonal matrix whose diagonal elements are $(\kappa(\lambda_j^k)^{\kappa-1} \mathbf{I}_S)_{1 \leq j \leq Q}$.

Updating $q(\mathbf{x})$

$$\begin{aligned}q^{k+1}(\mathbf{x}) &\propto \exp \left(\left\langle \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^k, \boldsymbol{\lambda}^k) \right\rangle_{q^k(\gamma)} \right) \\ &\propto \exp \left(\int \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^k, \boldsymbol{\lambda}^k) q^k(\gamma) d\gamma \right) \\ &= \mathcal{N}(\mathbf{x}; \mathbf{m}_{k+1}, \boldsymbol{\Sigma}_{k+1})\end{aligned}$$

$$\begin{cases} \boldsymbol{\Sigma}_{k+1}^{-1} &= \frac{1}{2} \delta^{-\frac{3}{2}} \mathbf{H}^\top \text{Diag}(\tilde{\mathbf{y}}) \mathbf{H} + 2 \mathbb{E}_{q^k(\gamma)} \mathbf{D}^\top \boldsymbol{\Lambda}^k \mathbf{D}, \\ \mathbf{m}_{k+1} &= \boldsymbol{\Sigma}_{k+1} \mathbf{H}^\top \mathbf{u}, \end{cases} \quad (1)$$

IMPLEMENTATION ISSUES:

- $\boldsymbol{\Sigma}_{k+1}$ is approximated by a diagonal matrix
- \mathbf{m}_{k+1} is approximated iteratively using conjugate gradient

Updating \mathbf{w} and λ

$$\begin{aligned}\mathbf{w}^{k+1} &= \underset{\mathbf{w}}{\operatorname{argmin}} \mathcal{KL}(q^{k+1}(\mathbf{x})q^k(\gamma) \| L(\Theta|\tilde{\mathbf{y}}; \mathbf{w}^{k+1}, \lambda)) \\ &= \underset{\mathbf{w}}{\operatorname{argmin}} \sum_{i=1}^M T_i(\tilde{y}_i, [\mathbf{H}\mathbf{m}_{k+1}]_i; w_i), \\ &= \max\{[\mathbf{H}\mathbf{m}_{k+1}]_i, 0\}\end{aligned}$$

$$\begin{aligned}\lambda_j^{k+1} &= \underset{\lambda_j \in [0, +\infty)}{\operatorname{argmin}} \mathcal{KL}(q^{k+1}(\mathbf{x})q^k(\gamma) \| L(\Theta|\tilde{\mathbf{y}}; \mathbf{w}^{k+1}, \lambda)). \\ &= \mathbb{E}_{q^{k+1}(\mathbf{x})} [\|\mathbf{D}_j\mathbf{x}\|^2] \\ &= \|\mathbf{D}_j\mathbf{m}_{k+1}\|^2 + \operatorname{trace} [\mathbf{D}_j^\top \mathbf{D}_j \boldsymbol{\Sigma}_{k+1}]\end{aligned}$$

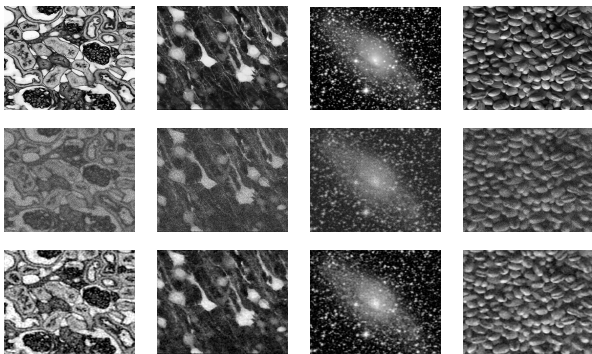
Updating $q(\gamma)$

$$\begin{aligned}q^{k+1}(\gamma) &\propto \exp \left(\left\langle \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^{k+1}, \boldsymbol{\lambda}^{k+1}) \right\rangle_{q^{k+1}(\mathbf{x})} \right) \\ &\propto \exp \left(\int \ln L(\mathbf{x}, \gamma, \mathbf{y}; \mathbf{w}^{k+1}, \boldsymbol{\lambda}^{k+1}) q^{k+1}(\mathbf{x}) d\mathbf{x} \right) \\ &= \mathcal{G}(\gamma; a_{k+1}, b_{k+1}) \\ &\begin{cases} a_{k+1} = \frac{N}{2\kappa} + \alpha = a, \\ b_{k+1} = \sum_{j=1}^Q (\lambda_j^{k+1})^\kappa + \beta, \end{cases}\end{aligned}$$

Experiments

- Different count levels x_{\max}
- Comparison with variational approaches:
 - ▶ different PG likelihood approximations: Generalized Anscombe Transform (GAST), the Exponential likelihood (Exp), the Exact likelihood (Exact)
 - ▶ different optimization algorithms: Spectral projected gradient [Bajic et al. 1994], Primal-dual splitting algorithm [Chouzenoux et al. 2015]

Visual results



From top to bottom: Original images: $x_{\max} = (20, 60, 100, 150)$.
Degraded images: $\text{SSIM} = (0.232, 0.423, 0.561, 0.586)$. Restored images
with the proposed approach: $\text{SSIM} = (0.575, 0.653, 0.765, 0.830)$.

Comparison with variational methods

		MAP (GAST) [PG]	MAP (GAST) [PD]	MAP(EXP) [PD]	MAP (Exact)[PD]	BV (GAST)
First image (350×350): $x_{\max} = 20$ h : Uniform 5×5 , $\sigma^2 = 9$	γ	fixed	fixed	fixed	fixed	automatic
	SNR	13.61	13.60	13.72	13.73	13.80
	Time (s.)	2897	490	3124	48587	29
Second image (257×256): $x_{\max} = 60$ h : Gaussian 9×9 , std 0.5, $\sigma^2 = 36$	γ	fixed	fixed	fixed	fixed	automatic
	SNR	15.35	15.33	15.42	15.43	15.22
	Time (s.)	3168	86	112	612	7
Third image (256×256): $x_{\max} = 100$ h : Uniform 3×3 , $\sigma^2 = 36$	γ	fixed	fixed	fixed	fixed	automatic
	SNR	13.71	13.77	13.81	13.81	14.17
	Time (s.)	2921	578	1060	17027	9
Fourth image (256×256): $x_{\max} = 150$ h : Gaussian 7×7 , std 1, $\sigma^2 = 40$	γ	fixed	fixed	fixed	fixed	automatic
	SNR	20.17	20.11	20.11	20.33	20.43
	Time (s.)	2964	886	3026	43397	14

PG: projected gradient algorithm

PD: primal dual algorithm

Simulations performed on an Intel(R) Xeon(R) CPU E5-2630, @ 2.40 GHz,
using a Matlab 7 implementation

- Comparable quantitative results with variational approaches
- Regularization parameter automatically tuned
- Competitive computation time

In a nutshell...

- ~> **Variational Bayesian method** for restoration of images corrupted with mixed Poisson Gaussian noise
- ~> Compared to variational approaches, the proposed method shows:
 - ✓ good quantitative and qualitative results
 - ✓ competitive computation time
- ~> Future work: extension to other PG likelihood approximations and others prior distributions