



THE UNIVERSITY of EDINBURGH
School of Engineering



Distributed Estimation of Latent Parameters in State Space Models Using Separable Likelihoods

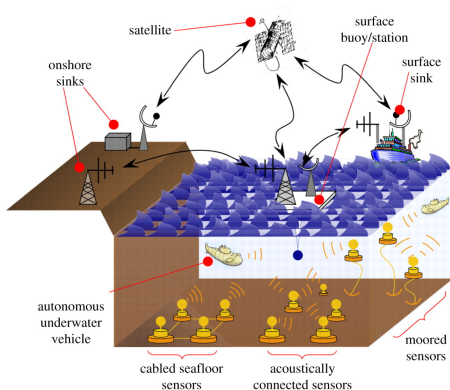
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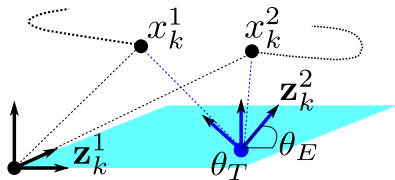
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Motivation



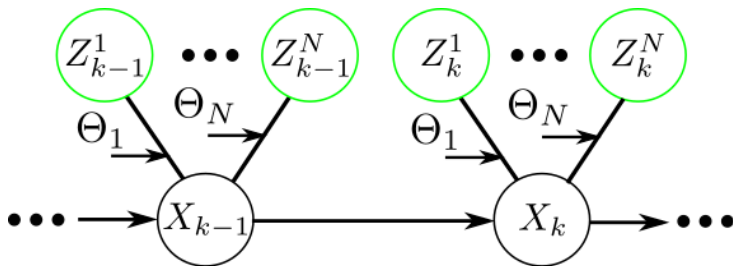
An underwater surveillance network (Heidemann et. al, 2011, adapted from Akyildiz, et.al, 2005)



Opportunistic sensor registration:
Estimation of respective sensor position θ_T and orientation θ_E using measurements \mathbf{z} from objects x over $k = 1, \dots, t$.
Example: Sensor localisation in GPS denying environments such as underwater.

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Multi-sensor state space model



- The object state X_k evolves as a Markov chain with transition density $\pi(x_k|x_{k-1})$ and initial density $\pi(x_1)$
- Sensor i measures z_k^i with a likelihood $l_i(z_k^i|x_k; \theta_i)$
- If latent parameters $\theta = [\theta_1, \dots, \theta_N]$ are known, only unknown is X_k which is estimated using sensor histories $z_{1:k}^1, \dots, z_{1:k}^N$
- Solved by finding $p(x_k|z_{1:k}^1, \dots, z_{1:k}^N, \theta)$ for $k = 1, \dots, t$ using Bayesian prediction and update recursions, i.e., “filtering”.

Centralised solution for estimating unknown θ

The likelihood for “parameter estimation in state space models”

$$l\left(z_{1:t}^1, \dots, z_{1:t}^N \mid \theta = [\theta_1, \dots, \theta_N]\right) = \prod_{k=1}^t p(z_k^1, \dots, z_k^N \mid z_{1:k-1}^1, \dots, z_{1:k-1}^N, \theta)$$

$$p(z_k^1, \dots, z_k^N \mid z_{1:k-1}^1, \dots, z_{1:k-1}^N, \theta) = \int \left(\prod_{j=1}^N l(z_k^j \mid x_k, \theta) \right) \times \underbrace{p(x_k \mid z_{1:k-1}^1, \dots, z_{1:k-1}^N, \theta)}_{\text{Prediction distribution of a (centralised) filter.}} d(x_k)$$

- The likelihood is a product of update terms over $k = 1, \dots, t$
- Computational cost is dominated by joint multi-sensor filtering.
- The filtering complexity is combinatorial with the number of sensors N (dimensionality of θ), when the state-space model is equipped with models tackling uncertainties in the number of objects, measurement-object associations etc.

Separable likelihoods

- Provide approximate models based on single sensor filtering, i.e., use local prediction $p(x_k | z_{1:k-1}^j)$ and update $p(x_k | z_{1:k}^j)$ for $j \in \mathcal{V}$.
 - provide scalability with the number of sensors
 - exploit recent advances in single sensor filtering algorithms
 - align well with distributed fusion architectures
- First, we consider a pair of sensors i and j .

Quad-term separable likelihood

Approximate $l(z_{1:t}^i, z_{1:t}^j | \theta)$ with $\prod_{k=1}^t q(z_k^i, z_k^j | z_{1:k-1}^i, z_{1:k-1}^j, \theta)$

$$q(z_k^i, z_k^j | z_{1:k-1}^i, z_{1:k-1}^j, \theta) \triangleq \frac{1}{\kappa_k(\theta)} \left(p(z_k^i | z_{1:k}^j, \theta) p(z_k^j | z_{1:k-1}^i, \theta) \right)^{1/2} \\ \times \left(p(z_k^j | z_{1:k}^i, \theta) p(z_k^i | z_{1:k-1}^j, \theta) \right)^{1/2}$$

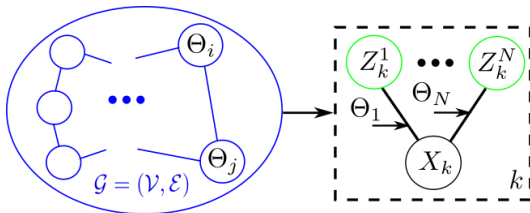
where $\kappa_k(\theta)$ is the normalisation constant.

Theorem (Kullback-leibler divergence of the approximation)

$$\begin{aligned}
 D(p||q) \leq & \frac{1}{2} \left(\left(H(X_k|Z_{1:k-1}^j, \Theta) - H(X_k|Z_{1:k-1}^i, Z_{1:k-1}^i, \Theta) \right) \right. \\
 & \left. + \left(H(X_k|Z_{k-1}^i, \Theta) - H(X_k|Z_{1:k-1}^j, Z_{1:k-1}^i, \Theta) \right) \right) \\
 & + \frac{1}{2} \left(\left(H(X_k|Z_{1:k}^j, \Theta) - H(X_k|Z_{1:k}^i, Z_{1:k-1}^i, \Theta) \right) \right. \\
 & \left. + \left(H(X_k|Z_{1:k}^i, \Theta) - H(X_k|Z_{1:k}^j, Z_{1:k-1}^i, \Theta) \right) \right), \quad (1)
 \end{aligned}$$

- Better approximation when local prediction and estimation at sensors i and j are accurate (small difference in Shannon Entropies H with respect to joint filtering).

Pairwise MRF Model with Separable Likelihood Edge Potentials



- Second, we assume that the local latent parameters Θ_i are Markov with respect to $\mathcal{G} = (\mathcal{V}, \mathcal{E})$ with edges associated with, e.g., available communication links, neighbourhood relations etc...

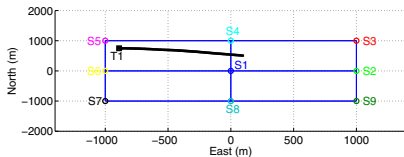
$$p(\theta | Z_{1:t}^1, \dots, Z_{1:t}^N) \propto \prod_{i \in \mathcal{V}} \psi_i(\theta_i) \prod_{(i,j) \in \mathcal{E}} \psi_{ij}^t(\theta_i, \theta_j),$$

$$\psi_i(\theta_i) = p_{0,i}(\theta_i), \quad \psi_{ij}^t(\theta_i, \theta_j) = \prod_{k=1}^t q(z_k^i, z_k^j | z_{1:k-1}^i, z_{1:k-1}^j, \theta_i, \theta_j)$$

Message passing algorithm for estimation of Θ

- 1: **for all** $j \in \mathcal{V}$ **do** ▷ Local filtering
- 2: Find $p(x_k | z_{1:k}^j)$ for $k = 1, \dots, t$
- 3: **end for**
- 4: **for all** $j \in \mathcal{V}$ **do** ▷ Sample from priors
- 5: Sample $\theta_j^{(l)} \sim p_{0,i}(\theta_j)$ for $l = 1, \dots, L$
- 6: **end for**
- 7: **for** $s = 1, \dots, S$ **do** ▷ S-steps of loopy belief propagation (LBP)
- 8: **for all** $(i, j) \in \mathcal{E}$ **do** ▷ Evaluate separable likelihood edge potentials
- 9: Find $\psi_{i,j}^t(\theta_i^{(l)}, \theta_j^{(l)}) = \prod_{k=1}^t q(z_k^i, z_k^j | z_{1:k-1}^i, z_{1:k-1}^j, \theta_i^{(l)}, \theta_j^{(l)})$ for $l = 1, \dots, L$
- 10: **end for**
- 11: **for all** $(i, j) \in \mathcal{E}$ **do** ▷ Find LBP messages m_{ji} s
- 12: Sample $\tilde{\theta}_i^{(l)}$ from $m_{ji}(\theta_i) = \int \psi_{ij}^t(\theta_i, \theta_j) \psi_j(\theta_j) \prod_{i' \in \text{ne}(j) \setminus i} m_{i'j}(\theta_j) d\theta_j$ for $l = 1, \dots, L$
- 13: **end for**
- 14: **for all** $i \in \mathcal{V}$ **do** ▷ Update local marginals $p_i(\theta_i)$ s
- 15: Sample $\theta_i^{(l)}$ from $p_i(\theta_i) \propto \psi_i(\theta_i) \prod_{j \in \text{ne}(i)} m_{ji}(\theta_i)$ for $l = 1, \dots, L$
- 16: $\hat{\theta}_i \leftarrow \frac{1}{L} \sum_{l=1}^L \theta_i^{(l)}$
- 17: **end for**
- 18: **end for**

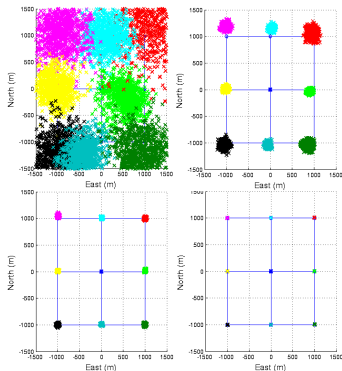
Example



- Linear model and additive Gaussian uncertainties
- θ_i s are unknown sensor locations
- sensor 1 is the origin of the network coordinate frame

$$\pi(\mathbf{x}_k | \mathbf{x}_{k-1}) = \mathcal{N}(\mathbf{x}_k; \mathbf{F}\mathbf{x}_{k-1}, \mathbf{Q})$$

$$l_i(z_k^i | \mathbf{x}_k; \theta_i) = \mathcal{N}(z_k^i; \mathbf{H}_i(\mathbf{x}_k - \theta_i), \mathbf{R}_i)$$



- $< 10m$ average error
- Average of 12.2s per edge per iteration (compared to 28.3s with joint filtering edge pot.)

Conclusion

- We consider multi-sensor state space models underpinning fusion networks and surveillance applications.
- We address scalability of latent parameter estimation in these models with the number of sensors.
- We propose a quad-term separable likelihood which together with pairwise MRFs facilitate scalability and distributed estimation.
- Previously, we introduced a dual-term separable likelihood [1], which can easily be used with random finite set models and solve problems involving multiple objects, measurements with false alarms, missed detections and association uncertainties.
- The proposed likelihood has a smaller error bound and can be used with hypothesis based multi-object filters [2]. It is not straightforward to adopt it for random finite set models, however.

[1] Uney, Mulgrew, Clark "A cooperative approach to sensor localisation in distributed fusion networks," IEEE TSP, March 2016.

[2] Uney, Mulgrew, Clark "Latent parameter estimation in fusion networks using separable likelihoods," IEEE TSIPN, submitted to the special issue on inference and learning over networks.