

# Cramer-Rao bound for the Time-Varying Poisson

Xinhui Rong and Victor Solo

School of Electrical Engineering & Telecommunications, UNSW, Sydney, Australia  
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## Problems & results:

Periodic Time-varying Poisson new results:

1. the asymptotic Cramer-Rao bound,
2. a maximum likelihood parameter estimation.

## Time-varying Poisson (tvP):

- A sequence of random event times with independent increment, observed in period  $[0, T]$ .

- Event times (ETs)  $T_1^n = \{T_1, \dots, T_n\}$ .

- Counting process  $N_t = N(0, t) = \#$  of events up to and including time  $t$ .

- (deterministic) tvP Intensity function:

$$\lambda(t) = \lim_{\delta \downarrow 0} \frac{1}{\delta} \Pr[N_{t+\delta} - N_t = 1],$$

- Log-likelihood [1]:

$$\mathcal{L}_T = \int_0^T \ln \lambda(t) dN_t - \int_0^T \lambda(u) du.$$

## Cramer-Rao bound (CRB) for tvP:

- $\theta$ : parameter vector of dimension  $d$ .

**Theorem 1** tvP CRB [1]. For an unbiased estimator  $\hat{\theta}$ ,

$$V_T = \text{var}[\hat{\theta}] \geq I_T^{-1}$$

where  $I_T$  is the Fisher Information Matrix (FIM)

$$I_T = \int_0^T \frac{1}{\lambda(t)} \frac{\partial \lambda(t)}{\partial \theta} \frac{\partial \lambda(t)}{\partial \theta^\top} dt,$$

and  $A \geq B$  means  $A - B$  is positive semi-definite.

## Model assumption:

- We assume periodic tvP intensity

$$\lambda(t) = b + a \cos(\omega t),$$

- So  $\theta = [b, a, \omega]^\top$ .

Preliminaries

Asymptotic CRB MLE algorithm

## Averages of periodic functions:

- Suppose  $\gamma(t) \geq 0$  has period  $\frac{2\pi}{\omega}$ .

- Then  $\kappa(\tau) = \gamma(\tau/\omega)$  has period  $2\pi$

**Result I** Consider the integral

$$\mathcal{J}_T = \frac{1}{T^{p+1}} \int_0^T t^p \gamma(t) [1 - \alpha f_{q,r}(\omega t)] dt,$$

where  $f_{q,r}(\tau) = \sin^q(\tau) \cos^r(\tau)$  and  $p, q, r \geq 0$  are integers and  $|\alpha| < 1$ . Then,

$$\mathcal{J}_T \rightarrow \frac{1}{p+1} L_{q,r}(\alpha), \text{ as } T \rightarrow \infty$$

$$L_{q,r}(\alpha) = \int_0^{2\pi} \kappa(\xi) [1 - \alpha f_{q,r}(\xi)] \frac{d\xi}{2\pi}$$

## Asymptotic CRB simplification:

- Introduce matrix  $D_T = \text{diag}(T^{\frac{1}{2}}, T^{\frac{1}{2}}, T^{\frac{3}{2}})$ .

**Result II** Asymptotic CRB.

$$\begin{aligned} D_T V_T D_T &\geq D_T I_T^{-1} D_T \\ &= (D_T^{-1} I_T D_T^{-1})^{-1} \rightarrow V_* \\ V_* &= \begin{bmatrix} bG & 0 \\ 0 & \frac{3R}{a\rho} \end{bmatrix} \& \rho = \frac{a}{b} \\ G &= \begin{bmatrix} 1 & \rho \\ \rho & R \end{bmatrix} \& R = 1 + \sqrt{1 - \rho^2} \end{aligned}$$

## Remarks.

$\text{var}[\hat{a}], \text{var}[\hat{b}]$  decay at rate  $T^{-1}$  but

$\text{var}[\hat{\omega}]$  decays at  $T^{-3}$ :

super efficiency of the frequency estimator.

## MLE algorithm via cyclic ascent:

- First MLE for periodic tvP, surprisingly.
- Reparametrization: 2 positive components
 
$$\begin{aligned} \lambda(t) &= b + a \cos(\omega t) \\ &= c + a[1 + \cos(\omega t)] \end{aligned}$$
- MLE via cyclic ascent:
  - Stage I: update  $\hat{c}$  and  $\hat{a}$  by EM
  - Stage II: update  $\hat{\omega}$  by Newton-Raphson.
- Formulae to be found in the paper.

## Scale-free(SF) parameters and measure:

- SF parameters:  $n_b = bT, m_T = \frac{\omega T}{2\pi}, \rho = \frac{a}{b}$

- SF precision measure:  $\pi(\hat{\theta}) = \frac{\theta}{se(\hat{\theta})}$ , where  $se(\hat{\theta})$  is the standard error of the estimator.

- SF precision expressed in SF parameters:

$$\pi(\hat{b}) = \sqrt{n_b}, \quad \pi(\hat{a}) = \sqrt{n_b} \sqrt{\gamma}, \quad \pi(\hat{\omega}) = \frac{m_T \sqrt{n_b} \sqrt{\gamma}}{2\pi \sqrt{3}}$$

where  $\sqrt{\gamma} = \frac{\rho}{\sqrt{R}}$ .

- Interpretations:

1. Precision only increases linearly with  $\sqrt{n_b}$ ;
2. But  $\hat{\omega}$  precision also rises linearly with  $m_T$ : greater precision;
3. If  $\rho \approx 0$ , precision for  $\hat{a}$  and  $\hat{\omega}$  very low.

## Simulation setup:

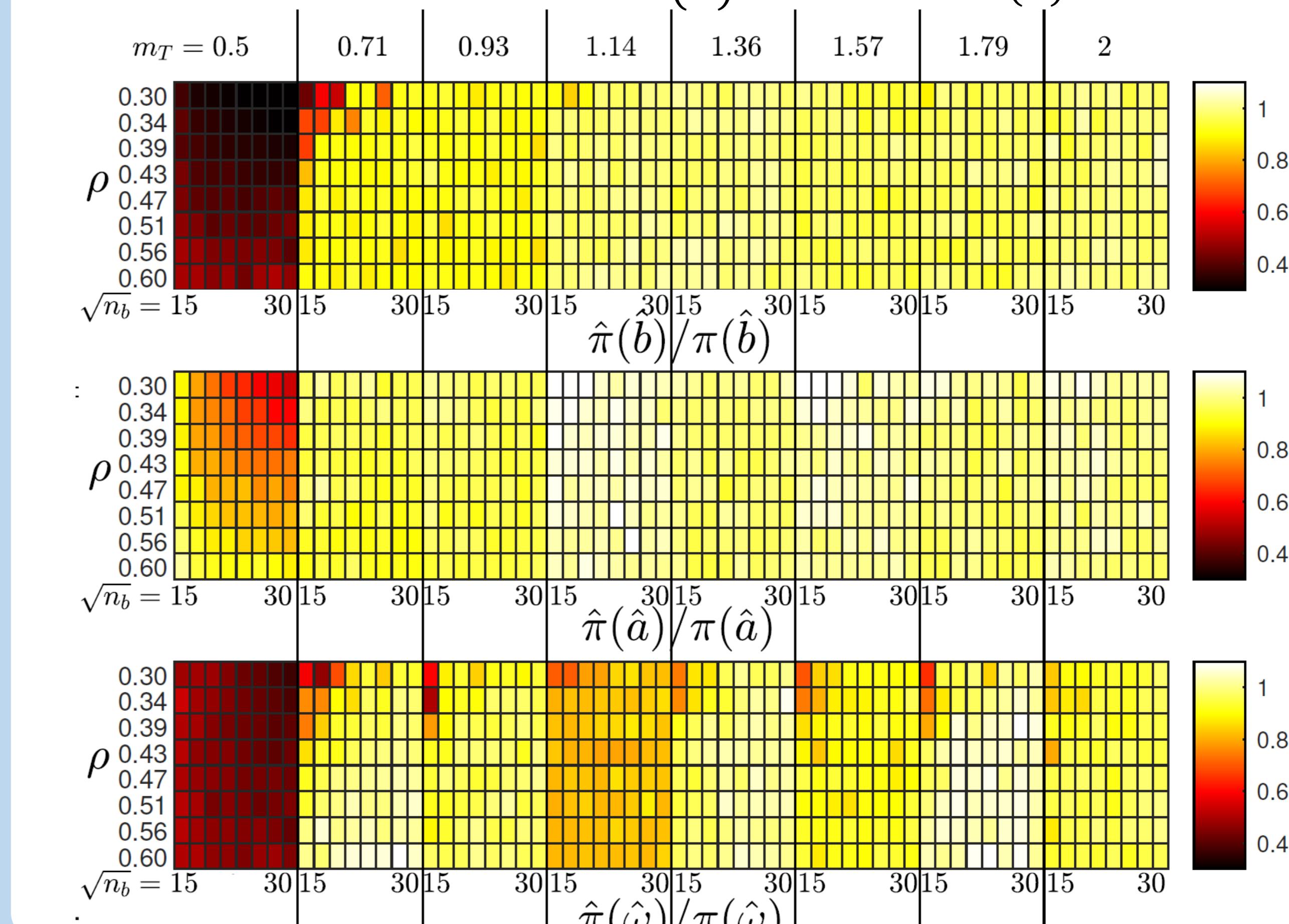
- Parameters to be chosen:  $b, a, \omega, T$ .

- One degree of freedom: fix  $\omega = 3$ .

- We vary SF parameters on a  $8 \times 8 \times 8$  grid:  $n_b = 15^2 : 30^2, m_T = 0.5 : 2, \rho = 0.3 : 0.6$

- Simulate 1,000 repeats on each grid point to estimate the precision:  $\hat{\pi}(\hat{\theta}) = \text{mean}(\hat{\theta})/\widehat{se}(\hat{\theta})$

**Result:** Heatmaps of  $\frac{\hat{\pi}(\hat{\theta})}{\pi(\hat{\theta})}$ . Expect  $\frac{\hat{\pi}(\hat{\theta})}{\pi(\hat{\theta})} \rightarrow 1$



## Reference:

[1] D. Snyder and M. Miller, Random Point Processes in Time and Space, Springer-Verlag, New York, 1991.