

DISTRIBUTED BAYESIAN TRACKING ON THE SPECIAL EUCLIDEAN GROUP USING LIE ALGEBRA PARAMETRIC APPROXIMATIONS

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1. Introduction

- The joint rotational and translational state of a rigid body can be parameterized as an element of the **Special Euclidean Group** $SE(3)$.
- Modern engineering systems involve cooperation between **multiple agents** on a **partially connected** network to run a common task, e.g., estimate a hidden state.
- In previous works, we introduced diffusion particle filters (PF) to perform **cooperative tracking** of states that evolved on the Spherical and the Stiefel manifolds and the **Special Orthogonal Group**.
- Diffusion PFs include a **data assimilation step** where agents update their beliefs about the unknown states, assimilating **local** measurements and measurements from **neighboring nodes**.
- The local updated beliefs are then **exchanged** between nodes in a **compressed** form, using Gaussian parametric approximations on the **Lie Algebra** associated to $SE(n)$.

2. Special Euclidean Group $SE(n)$

- The **Special Euclidean Group** $SE(n)$ is a **matrix Lie group**. An element S of $SE(n)$ is given as

$$S = \begin{bmatrix} \Omega & \mathbf{u} \\ \mathbf{0}_{1 \times n} & 1 \end{bmatrix},$$

- where Ω is a member of the **Special Orthogonal Group** $SO(n)$, $\mathbf{u} \in \mathbb{R}^n$, and $\mathbf{0}_{1 \times n}$ denotes a vector with null entries.
- The group $SE(n)$ has dimension $d \triangleq n(n+1)/2$ and, for $n=3$, it corresponds to the set of all possible **translations and rotations** of a 3-dimensional rigid object.
- $SE(n)$ is also a **differentiable manifold**. Thus, we can define a **tangent space** \mathcal{T}_S at each point $S \in SE(n)$.
- The **Lie algebra** $\mathfrak{se}(n)$ is, by definition, the tangent space to the identity matrix I , i.e., \mathcal{T}_I .
- A matrix $S \in SE(n)$ can be mapped into a matrix $X \in \mathfrak{se}(n)$ using the **logarithmic map** $\text{Log} : SE(n) \rightarrow \mathfrak{se}(n)$.
- For $SE(n)$, $\text{Log}(S)$, is the usual **matrix logarithmic function**, denoted logm .
- For a matrix $S \in SE(n)$,

$$\text{Log}(S) = \begin{bmatrix} \text{logm}(\Omega) & \mathbf{V}\mathbf{u} \\ \mathbf{0}_{1 \times n} & 0 \end{bmatrix},$$

- where $\mathbf{V} = \mathbf{I} + \sum_{m=1}^{\infty} \Omega^m / (m+1)!$.
- For $n=3$, $\mathbf{Z} = \text{logm}(\Omega)$ has the form

$$\mathbf{Z} = \begin{bmatrix} 0 & -c & b \\ c & 0 & -a \\ -b & a & 0 \end{bmatrix}.$$

- Thus, $X = \text{Log}(S) \in \mathfrak{se}(n)$ is **isomorphic** to a vector $\mathbf{x} \in \mathbb{R}^d$, and we can define a **bijjective mapping** $\Phi : \mathfrak{se}(n) \rightarrow \mathbb{R}^d$, such that $\Phi(X) = [z^T (\mathbf{V}\mathbf{u})^T]^T \in \mathbb{R}^d$ where z is a vector that collects the free-varying entries of \mathbf{Z} .

3. Problem Setup

- Let $S_k \in SE(n)$ denote an unknown **state** at time $k \geq 0$ that evolves according to the **random walk**

$$S_k = S_{k-1} \text{Exp}(\Phi^{-1}(\epsilon_k)), \quad k > 0,$$

with $p(S_0) \propto 1$, where $\{\epsilon_k\}$ is a sequence of i.i.d. Gaussian random vectors in \mathbb{R}^d with zero mean and covariance matrix Λ_k .

- The nodes record at each instant k the **observations**

$$\mathbf{Y}_{k,r} = \mathcal{H}_r(\Pi S_k) + \mathbf{W}_{k,r}, \quad k > 0,$$

where $r \in \{1, \dots, R\}$ denotes the r -th node in the network, $\mathcal{H}_r(\cdot) : \mathbb{R}^{n \times (n+1)} \mapsto \mathbb{R}^{n \times (n+1)}$ is a general function, $\Pi \in \mathbb{R}^{n \times (n+1)}$ is an $(n+1) \times (n+1)$ identity matrix without its bottommost row, and $\{\mathbf{W}_{k,r}\}$ is a sequence of i.i.d. samples of a **Matrix Gaussian p.d.f.** $\mathcal{N}_{n,n+1}(\mathbf{W}_{k,r} | \mathbf{0}_{n,n+1}, \Psi_r, \Gamma_r)$.

- Given the observations $\{\mathbf{Y}_{l,r}\}$, $0 \leq l \leq k$, $0 \leq r \leq R$, our goal is to **recursively estimate** S_k in a **distributed** fashion.

4. RndEx Diffusion PF

- The **RndEx Algorithm** has two steps: **Random Exchange** and **Data Assimilation**.

- In the **Random Exchange** step, a node l exchanges with another randomly chosen node r its posterior p.d.f. $p(S_{k-1} | \tilde{\mathbf{Y}}_{1:k-1,l})$, in which $\tilde{\mathbf{Y}}_{1:k-1,l}$ denotes all observations assimilated up to instant k .

- Suppose that the posterior p.d.f. is approximated by the weighted particle set $\{w_{k-1,l}^{(q)}, S_{k-1,l}^{(q)}\}$, $q = 1, \dots, Q$, $Q \gg 1$.

- Before the exchange, node l **compresses** the representation using a **Gaussian parametric approximation** as follows.

- Given $\{w_{k-1,l}^{(q)}, S_{k-1,l}^{(q)}\}$ compute its **centroid** $\hat{S}_{k-1,r}$ as

$$\hat{S}_{k-1,l}^{<i+1>} = \hat{S}_{k-1,l}^{<i>} \cdot \expm \left[\sum_{q=1}^Q w_{k-1,l}^{(q)} \logm \left((\hat{S}_{k-1,l}^{<i>})^{-1} S_{k-1,l}^{(q)} \right) \right],$$

where $\hat{S}_{k-1,l}^{<i>}$ denotes the i -th estimate of the weighted average, with $\hat{S}_{k-1,l}^{<0>}$ chosen as a random element of the particle set.

- Then, we evaluate $\mathbf{x}_{k-1,l}^{(q)} = \Phi \left(\logm \left(\hat{S}_{k-1,l}^{-1} S_{k-1,l}^{(q)} \right) \right) \in \mathbb{R}^d$.

- The weighted particle set $\{w_{k-1,l}^{(q)}, \mathbf{x}_{k-1,l}^{(q)}\}$ is then approximated by a **Gaussian** p.d.f. via **moment matching**, with sample moments $\bar{\mathbf{x}}_{k-1,l}$ and $\Sigma_{k-1,l}$.

- At the end of the Random Exchange Step, node r receives $\{\bar{\mathbf{x}}_{k-1,l}, \Sigma_{k-1,l}, \hat{S}_{k-1,l}\}$ and **rebuilds** the particle set as

$$\begin{aligned} \bar{\mathbf{x}}_{k-1,l}^{(q)} &\sim \mathcal{N}(\mathbf{x}_{k-1,l} | \bar{\mathbf{x}}_{k-1,l}, \Sigma_{k-1,l}), \\ \hat{S}_{k-1,l}^{(q)} &= \hat{S}_{k-1,l} \expm \left(\Phi^{-1}(\bar{\mathbf{x}}_{k-1,l}^{(q)}) \right) \\ \bar{w}_{k-1,l}^{(q)} &= \frac{1}{Q}, \end{aligned}$$

where $q \in \{1, \dots, Q\}$ and $\mathcal{N}(s | \mu; \Theta)$ denotes a (vector) multivariate Gaussian p.d.f. with argument s , parameterized by the mean vector μ and the covariance matrix Θ .

- In the **Data Assimilation** step, node r samples new particles from the prior importance function

$$S_{k,r}^{(q)} \sim p(S_k | S_{k-1,l}^{(q)}),$$

and updates the particles' weights as

$$w_{k,r}^{(q)} \propto w_{k-1,l}^{(q)} \left[\prod_{u \in \tilde{N}(r)} p(\mathbf{Y}_{k,u} | S_{k,r}^{(q)}) \right],$$

in which $\tilde{N}(r)$ denotes the r -th node closed neighborhood.

- The updated set $\{w_{k,r}^{(q)}, S_{k,r}^{(q)}\}$ is then a Monte Carlo representation of the **posterior p.d.f.** $p(S_k | \tilde{\mathbf{Y}}_{k,r}, \tilde{\mathbf{Y}}_{1:k-1,l}) \triangleq p(S_k | \tilde{\mathbf{Y}}_{1:k,r})$.

- State estimates** can be computed as the centroid of this particle set.

5. Alternative Formulation

- Alternatively**, the state components $\Omega_{k-1,l}^{(q)}$ and $\mathbf{u}_{k-1,l}^{(q)}$ can be separately compressed:

- Given $\{w_{k-1,l}^{(q)}, S_{k-1,l}^{(q)}\}$ compute the **centroids**

$$\hat{\Omega}_{k-1,l}^{<i+1>} = \hat{\Omega}_{k-1,l}^{<i>} \expm \left[\sum_{q=1}^Q w_{k-1,l}^{(q)} \logm \left((\hat{\Omega}_{k-1,l}^{<i>})^{-1} \Omega_{k-1,l}^{(q)} \right) \right],$$

$$\hat{\mathbf{u}}_{k-1,l} = \sum_{q=1}^Q w_{k-1,l}^{(q)} \mathbf{u}_{k-1,l}^{(q)},$$

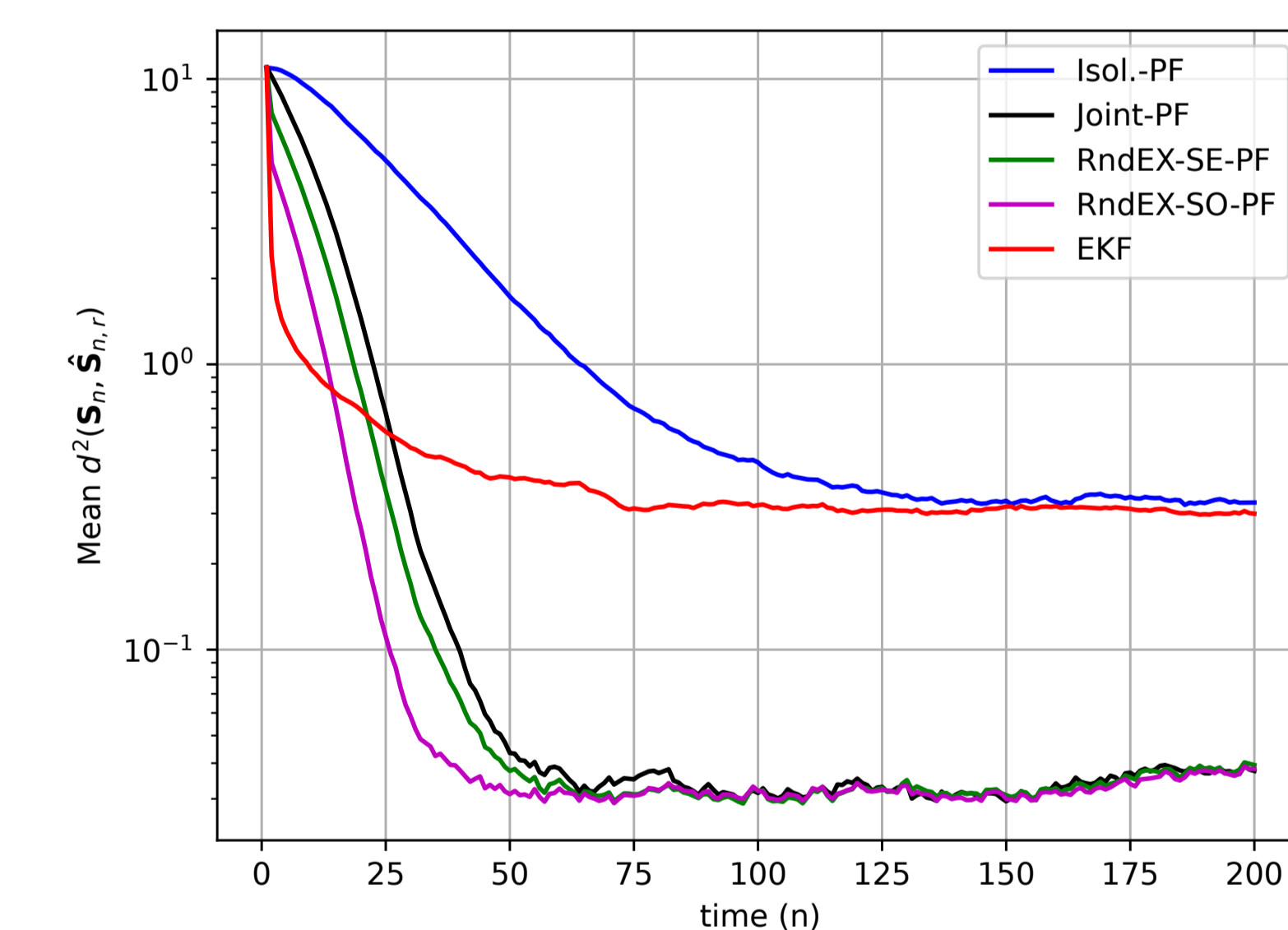
- Then, evaluate $\mathbf{x}_{k-1,l}^{(q)} = \Phi \left(\logm \left(\hat{\Omega}_{k-1,l}^{-1} \Omega_{k-1,l}^{(q)} \right) \right) \in \mathbb{R}^d$.

- The weighted particle set $\{w_{k-1,l}^{(q)}, \mathbf{x}_{k-1,l}^{(q)}, \mathbf{u}_{k-1,l}^{(q)}\}$ is then approximated by a **Gaussian** p.d.f. via **moment matching**.

- At the end of the Random Exchange Step, node r receives $\{\bar{\mathbf{x}}_{k-1,l}, \bar{\mathbf{u}}_{k-1,l}, \Sigma_{k-1,l}, \hat{\Omega}_{k-1,l}\}$ and **rebuilds** the particle set as previously.

6. Simulation Results

- We performed a **numerical simulation** with 300 independent trials.
- We set $n=3$ and $Q=300$. In each trial, 200 synthetic data samples were generated.
- We used a network with five nodes: nodes 1 to 4 are on the vertices of a square and node 5 is at its center and is connected to all other nodes.
- The **noise covariance** matrices were set to $\Psi_r = \mathbf{I}$ and $\Gamma_r = \mathbf{I} \cdot 10^{-\alpha_r/10}$, with α_r equal to 3, 6, 10, 13 and 20 dB for $r = 1, \dots, 5$, respectively. The **driving noise covariance** matrix Λ_k was set to $0.05 \mathbf{I}$.
- For **comparison**, we ran in the same setup three competing algorithms: i) bootstrap PFs operating isolatedly in each node (Isol.-PF), ii) a bootstrap PF with access to all observations (Joint-PF), and iii) a $SE(n)$ -constrained EKF.
- We assumed that $[\mathcal{H}_r(\mathbf{X})]_{i,j} = h([\mathbf{X}]_{i,j})$, and set $h(x) = \tanh(x/\beta)/\beta$ with $\beta = 1.7$.
- The algorithm's performance was evaluated in terms of the **squared geodesic distance** in $SE(n)$, i.e., $d^2(S_k, \hat{S}_k) = \|\logm(\Omega_k^T \hat{\Omega}_k)\|_F^2 + \|\mathbf{u}_k - \hat{\mathbf{u}}_k\|^2$.



7. Conclusions

- This paper described **two novel RndEx particle filters** for tracking the state of a dynamic system that evolves on the Special Euclidean Group.
- The first algorithm exchanges **Gaussian** parametric approximations built on a **vector space isomorphic** to $\mathfrak{se}(n)$.
- The second algorithm directly computes approximations for the translational information $\mathbf{u}_{k-1,l}^{(q)}$, performing computations on the Lie algebra of $\mathfrak{so}(n)$ **only for rotation matrices** $\Omega_{k-1,l}^{(q)}$.
- Experimental results show that the proposed methods perform similarly to a centralized PF-based estimator, greatly outperforming PFs operating in isolation and an extended Kalman filter, at increased computational cost.