

Multi-Sensor Multi-Scan Radar Sensing of Multiple Extended Targets

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Table of Contents

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

Martin Voigt Vejling

3

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

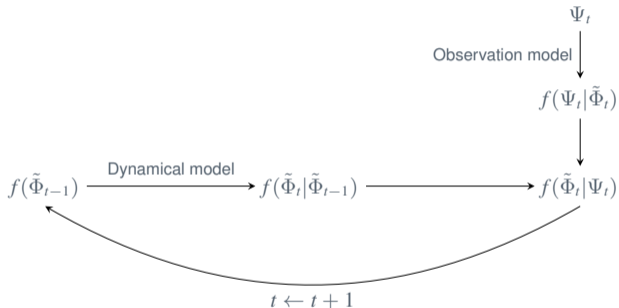


Figure: Typical tracking filter.

Martin Voigt Vejling

4

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

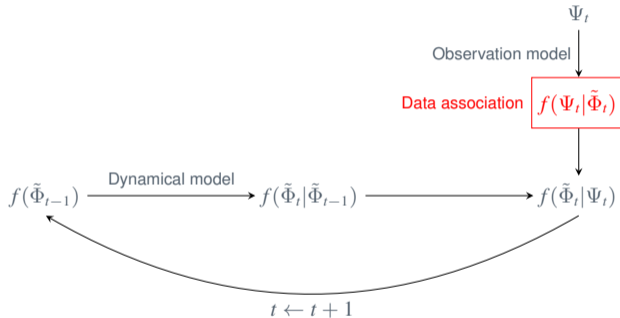


Figure: Typical tracking filter.

Martin Voigt Vejling

5

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

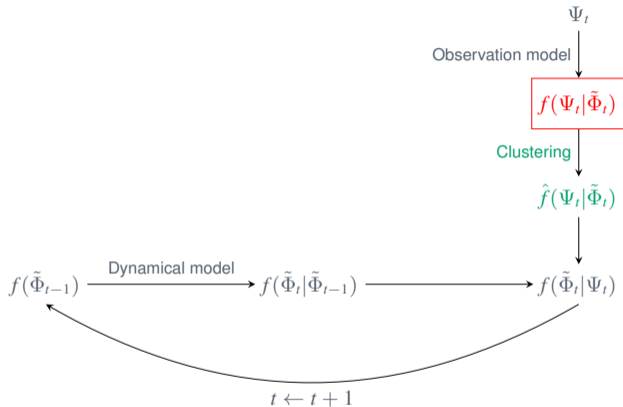


Figure: Typical filter for tracking multiple extended targets: clustering is used to only consider a subset of data association hypotheses in the likelihood.

¹Karl Granström, Maryam Fatemi, and Lennart Svensson. *Poisson Multi-Bernoulli Mixture Conjugate Prior for Multiple Extended Target Filtering*. IEEE Trans. Aerosp. Electron. Syst., 56(1):208-225, 2020

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

6

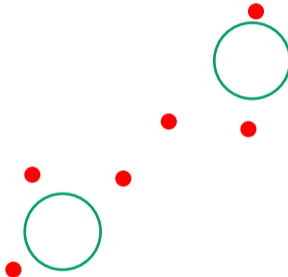


Figure: Observation.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

7

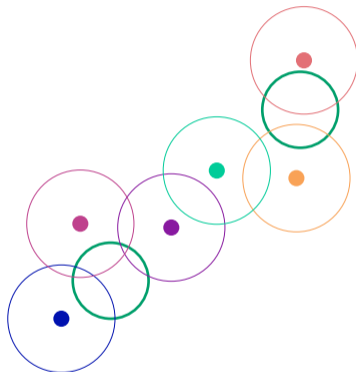


Figure: Distance based clustering: low radius.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

8

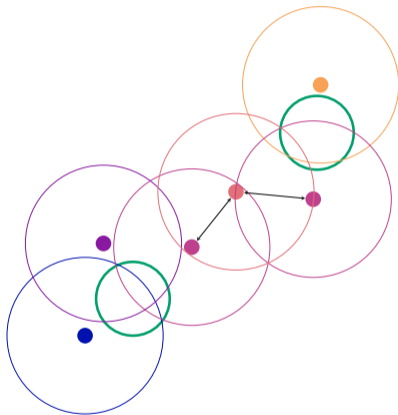


Figure: Distance based clustering: medium radius.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

9

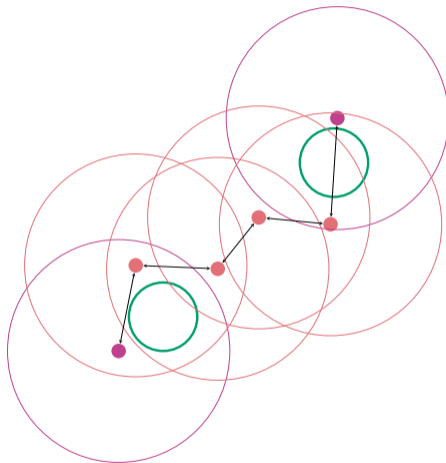


Figure: Distance based clustering: large radius.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

10

Scenario

Modelling details

Results

Conclusion

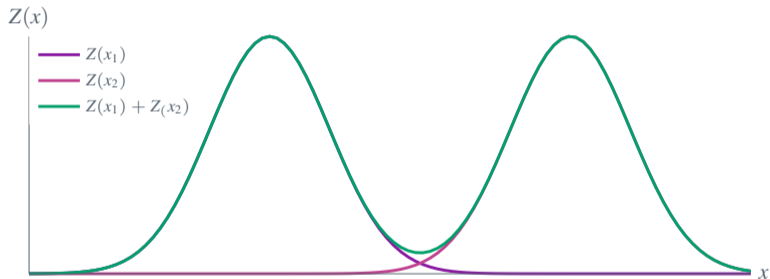


Figure: Well-separated targets.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

11

Scenario

Modelling details

Results

Conclusion



Figure: Separable targets.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

12

Scenario

Modelling details

Results

Conclusion

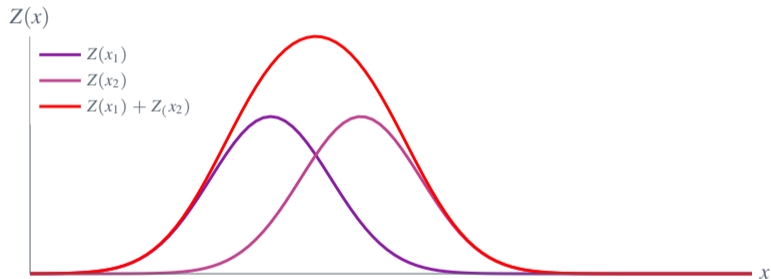


Figure: Non-separable targets.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

13

Modelling details

Results

Conclusion

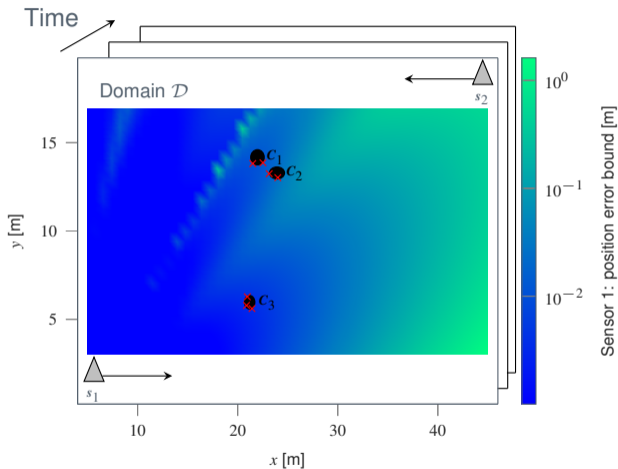


Figure: Illustration of scenario with two sensors and three extended targets.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

14

Modelling details

Results

Conclusion

► Targets: $\tilde{\Phi} = \{C_1, \dots, C_L\}$ for $C_l = (c_l, E_l)$.

► Measurement model: $z_{k,i}(C) = v_i + \varepsilon_{k,i}(v_i)$, $v_i = c + \varepsilon_i^E$.

► Measurement process: $\Psi_k = \{z_{k,i}(C_l)\}_{l,i} \cup \Psi_k^c$.

► Aggregated observations: $\Psi = \bigcup_{k=1}^K \Psi_k$.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

15

Assume $\Psi_k | \tilde{\Phi}$ is Poisson with intensity

$$Z_{\tilde{\Phi}, \lambda^c}^{(k)}(\boldsymbol{\xi}) = \lambda^c + \sum_{l=1}^L \eta_k(\boldsymbol{\xi} | \mathbf{C}_l) \quad (1)$$

where $\eta_k(\boldsymbol{\xi} | \mathbf{C}_l) = \tilde{\rho}_k(\mathbf{C}_l) \mathcal{K}_{\tilde{\Sigma}_k(\mathbf{C}_l)}(\boldsymbol{\xi} - \mathbf{c}_l)$ and

$$\mathcal{K}_{\Sigma}(\boldsymbol{\xi} - \mathbf{c}) = \frac{\exp\left(-\frac{1}{2}(\boldsymbol{\xi} - \mathbf{c})^T \Sigma^{-1}(\boldsymbol{\xi} - \mathbf{c})\right)}{\sqrt{(2\pi)^d |\Sigma|}}. \quad (2)$$

We call this a doubly inhomogeneous-generalized shot noise Cox process (DI-GSNCP).

21

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

16

The model parameters are $\theta = (\tilde{\Phi}, \lambda, \lambda^c)$ and, up to a constant, the posterior is $\Pi(\theta|\Psi) = \prod_{k=1}^K \Pi_{\text{likelihood}}(\Psi_k|\tilde{\Phi}, \lambda^c) \Pi_{\text{prior}}(\tilde{\Phi}, \lambda, \lambda^c)$. Using the Poisson assumption, the observed likelihood is

$$\Pi_{\text{likelihood}}(\Psi_k|\tilde{\Phi}, \lambda^c) = \exp((1 - \lambda^c)|\mathcal{D}| - \sum_{l=1}^L \tilde{\rho}_k(\mathbf{c}_l)) \prod_{m_k=1}^{M_k} Z_{\tilde{\Phi}, \lambda^c}^{(k)}(\mathbf{p}_{k,m_k}). \quad (3)$$

The prior is proportional to

$$\Pi_{\text{prior}}(\tilde{\Phi}, \lambda, \lambda^c) = \lambda^L \mathbb{1}_{\mathbb{R}_+}(\lambda) \mathbb{1}_{\mathbb{R}_+}(\lambda^c) \prod_{l=1}^L \Pi_E(\Sigma_l^E) \prod_{\substack{j=1 \\ j \neq l}}^L \mathbb{1}_{\mathbb{R}_+}(\|\mathbf{c}_l - \mathbf{c}_j\| - R) \quad (4)$$

where $\mathbb{1}_A(x)$ is the indicator function which equals 1 if $x \in A$ and 0 otherwise, Π_E is the extent prior density, and we have assumed uniform priors on the cluster center and clutter intensities.

21

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

17

Results

Conclusion

- ▶ Using Markov chain Monte Carlo (MCMC), a Markov chain $\theta_0, \theta_1, \dots$ is constructed as follows: given the previous configuration in the Markov chain, i.e., θ_{i-1} , a possible new configuration of model parameters is sampled as $\theta^* \sim \mathcal{Q}_{(\theta_{i-1})}(\cdot)$ from a predefined transition density. We define the acceptance probability function

$$\alpha(\theta_{i-1}, \theta^*) = \min\left\{1, \frac{\Pi(\theta^*|\Psi)}{\Pi(\theta_{i-1}|\Psi)}\right\}$$

and sample $U \sim \text{Unif}(0, 1)$. If $U < \alpha(\theta_{i-1}, \theta^*)$ set $\theta_i = \theta^*$, otherwise let $\theta_i = \theta_{i-1}$.

- ▶ In this work, the parameters in θ are updated sequentially. Firstly, the driving process $\tilde{\Phi}$ is updated followed by an update of λ and λ^c . We define the transition density for updating the driving process as a birth-death-move proposal: with probability p_m a move is taken, with probability $(1 - p_m)p_b(\theta_{i-1})$ a birth is proposed, and with probability $(1 - p_m)(1 - p_b(\theta_{i-1}))$ a death is proposed.

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

18

- ▶ We define a scenario with two sensors scanning over 6 time epochs. The spatial domain is $\mathcal{D} = [0, 50] \times [0, 20] \times [0, 10] m^3$ and we assume extent prior as $e_{l,1}, e_{l,2}, e_{l,3} \sim \text{Unif}(1, 1.5)$ and $e_{l,4}, e_{l,5}, e_{l,6} \sim \text{Unif}(-0.5, 0.5)$. Moreover, we let $R = 8 m$, $\lambda = \frac{20}{|\mathcal{D}|}$, and $\lambda^c = \frac{35}{|\mathcal{D}|}$.
- ▶ Baselines: Oracle and DBSCAN.
- ▶ Performance metric: Optimal sub-pattern assignment (OSPA) metric with pair-wise metric defined by the Gaussian Wasserstein distance.

21

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

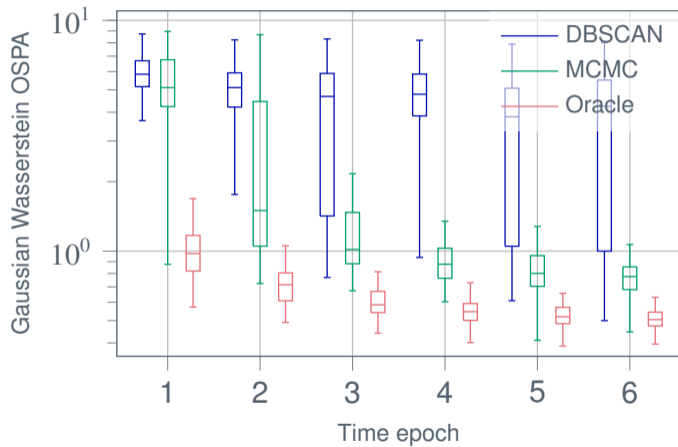
Scenario

Modelling details

Results

Conclusion

19



21

Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

20

- ▶ A methodology for target state estimation in multi-sensor multi-scan multiple extended target sensing scenarios is developed. The method is based on parametrizing the target states through a DI-GSNCP taking spatial properties of multiple sensors into account and using a model jump MCMC algorithm to estimate the parameters.
- ▶ The method scales only **linearly** in the number of measurements, effectively estimating the target states without requiring data association.
- ▶ Numerical experiments demonstrate the benefits over spatial proximity based clustering in **high clutter** scenarios with **closely spaced targets**.
- ▶ In future work, we aim to generalize the method to: (i) non-linear measurement models, (ii) moving targets, (iii) inter-cluster interactions.

21



Martin Voigt Vejling

Motivation

Tracking filter

Clustering example

Separation of clusters

Scenario

Modelling details

Results

Conclusion

Thank you for listening!

Questions?

21

21