

PLUG-IN MEASURE-TRANSFORMED QUASI-LIKELIHOOD RATIO TEST FOR RANDOM SIGNAL DETECTION



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PROBLEM FORMULATION

Detection of a random signal that lies on a known rank-one subspace:

$$H_0 : \begin{cases} \mathbf{X}_n = \mathbf{W}_n, & n = 1, \dots, N \\ \mathbf{Y}_m = \mathbf{W}_m^{(s)}, & m = 1, \dots, M \end{cases}$$

$$H_1 : \begin{cases} \mathbf{X}_n = S_n \mathbf{a} + \mathbf{W}_n, & n = 1, \dots, N \\ \mathbf{Y}_m = \mathbf{W}_m^{(s)}, & m = 1, \dots, M \end{cases}$$

GAUSS-GAUSS DETECTOR

- ▶ GLRT detector that assumes jointly Gaussian signal and noise.
- ▶ Advantages: simple implementation, ease of performance analysis.
- ▶ Disadvantage: sensitive to model mismatch.

PLUG-IN NSDD-GLRT

- ▶ Conditional GLRT detector that assumes a compound-Gaussian noise.
- ▶ The scatter matrix is replaced by noise-only secondary data ML estimate.
- ▶ Advantages: robust against heavy-tailed noise outliers.
- ▶ Disadvantage: Computationally demanding in high-dimensions, does not reject large-norm outliers.

MEASURE TRANSFORMED (MT) QGLRT: BASIC IDEA

- ▶ Selects a Gaussian probability model that best empirically fits a transformed probability measure of the data.
- ▶ By proper choice of the transform the MT-QGLRT can gain enhanced robustness to outliers.
- ▶ Have the computational and implementation advantages of the GGD.

PROBABILITY MEASURE TRANSFORM

Let $\mathbf{X} \in \mathbb{C}^p$. Define the measure space (\mathcal{X}, S_x, P_x) . Given a non-negative function $u : \mathbb{C}^p \rightarrow \mathbb{R}_+$ satisfying $0 < E[u(\mathbf{X}); P_x] < \infty$. A transform $\mathbb{T}_u : P_x \rightarrow Q_x^{(u)}$ is defined as:

$$\mathbb{T}_u[P_x](A) = Q_x^{(u)}(A) \triangleq \int_A \varphi_u(\mathbf{x}) dP_x(\mathbf{x}),$$

where $A \in S_x$, and

$$\varphi_u(\mathbf{x}) \triangleq \frac{u(\mathbf{x})}{E[u(\mathbf{X}); P_x]}.$$

The function $u(\cdot)$ is called the **MT-function**.

THE MEASURE-TRANSFORMED MEAN AND COVARIANCE

- ▶ MT-mean: $\boldsymbol{\mu}_x^{(u)} \triangleq E[\mathbf{X}\varphi_u(\mathbf{X}); P_x]$
- ▶ MT-covariance: $\boldsymbol{\Sigma}_x^{(u)} \triangleq E[\mathbf{X}\mathbf{X}^H\varphi_u(\mathbf{X}); P_x] - \boldsymbol{\mu}_x^{(u)}\boldsymbol{\mu}_x^{(u)H}$
- ▶ The mean and covariance under $Q_x^{(u)}$ can be estimated using samples from P_x .
 $\hat{\boldsymbol{\mu}}_x^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \hat{\varphi}_u(\mathbf{X}_n)$ and $\hat{\boldsymbol{\Sigma}}_x^{(u)} \triangleq \sum_{n=1}^N \mathbf{X}_n \mathbf{X}_n^H \hat{\varphi}_u(\mathbf{X}_n) - \hat{\boldsymbol{\mu}}_x^{(u)} \hat{\boldsymbol{\mu}}_x^{(u)H}$
 where $\hat{\varphi}_u(\mathbf{X}_n) \triangleq u(\mathbf{X}_n) / \sum_{n=1}^N u(\mathbf{X}_n)$

ROBUSTNESS TO OUTLIERS

- ▶ Define the ϵ -contaminated probability measure:

$$P_\epsilon \triangleq (1 - \epsilon)P_x + \epsilon\delta_y,$$

where $0 \leq \epsilon \leq 1$, $\mathbf{y} \in \mathbb{C}^p$, and δ_y is the Dirac probability measure at \mathbf{y} .

- ▶ The influence function (IF) [1] of an estimator with statistical functional $H[\cdot]$:

$$\text{IF}_H(\mathbf{y}; P_x) \triangleq \lim_{\epsilon \rightarrow 0} \frac{H[P_\epsilon] - H[P_x]}{\epsilon} = \left. \frac{dH[P_\epsilon]}{d\epsilon} \right|_{\epsilon=0}$$

- ▶ Describes the effect on the estimator of an infinitesimal contamination at \mathbf{y} .
- ▶ An estimator is said to be B-robust if its IF is bounded.

Proposition (Boundedness of the influence function)

The influence functions of the empirical MT-mean and MT-covariance are bounded if the MT-function $u(\mathbf{y})$ and the product $u(\mathbf{y})\|\mathbf{y}\|^2$ are bounded.

MEASURE-TRANSFORMED (MT) QGLRT: DERIVATION OF THE TEST

- ▶ Compares the empirical KLDs between $Q_x^{(u)}$ and two normal distributions that are characterized by the MT-mean and MT-covariance under each hypothesis:

$$T_u \triangleq \hat{D}_{\text{KL}}[Q_x^{(u)} || \Phi_{\mathbf{x}; H_0}^{(u)}] - \hat{D}_{\text{KL}}[Q_x^{(u)} || \Phi_{\mathbf{x}; H_1}^{(u)}]$$

$$= \left(D_{\text{LD}}[\hat{\boldsymbol{\Sigma}}_x^{(u)} || \boldsymbol{\Sigma}_{\mathbf{x}; H_0}^{(u)}] + \left\| \hat{\boldsymbol{\mu}}_x^{(u)} - \boldsymbol{\mu}_{\mathbf{x}; H_0}^{(u)} \right\|_{(\boldsymbol{\Sigma}_{\mathbf{x}; H_0}^{(u)})^{-1}}^2 \right)_{H_0}^{H_1}$$

$$- \left(D_{\text{LD}}[\hat{\boldsymbol{\Sigma}}_x^{(u)} || \boldsymbol{\Sigma}_{\mathbf{x}; H_1}^{(u)}] + \left\| \hat{\boldsymbol{\mu}}_x^{(u)} - \boldsymbol{\mu}_{\mathbf{x}; H_1}^{(u)} \right\|_{(\boldsymbol{\Sigma}_{\mathbf{x}; H_1}^{(u)})^{-1}}^2 \right)_{H_0}^{H_1} \stackrel{H_1}{\leq} \tau$$

- ▶ Equivalent test-statistic under any MT-function $u(\mathbf{x}) = v(\mathbf{P}_a^\perp \mathbf{x})$, $v : \mathbb{C}^p \rightarrow \mathbb{R}_+$

$$T_u' = \frac{\mathbf{a}^H (\boldsymbol{\Sigma}_w^{(u)})^{-1} \hat{\mathbf{C}}_x^{(u)} (\boldsymbol{\Sigma}_w^{(u)})^{-1} \mathbf{a}}{\mathbf{a}^H (\boldsymbol{\Sigma}_w^{(u)})^{-1} \mathbf{a}},$$

where $\hat{\mathbf{C}}_x^{(u)} \triangleq \hat{\boldsymbol{\Sigma}}_x^{(u)} + \hat{\boldsymbol{\mu}}_x^{(u)} \hat{\boldsymbol{\mu}}_x^{(u)H}$.

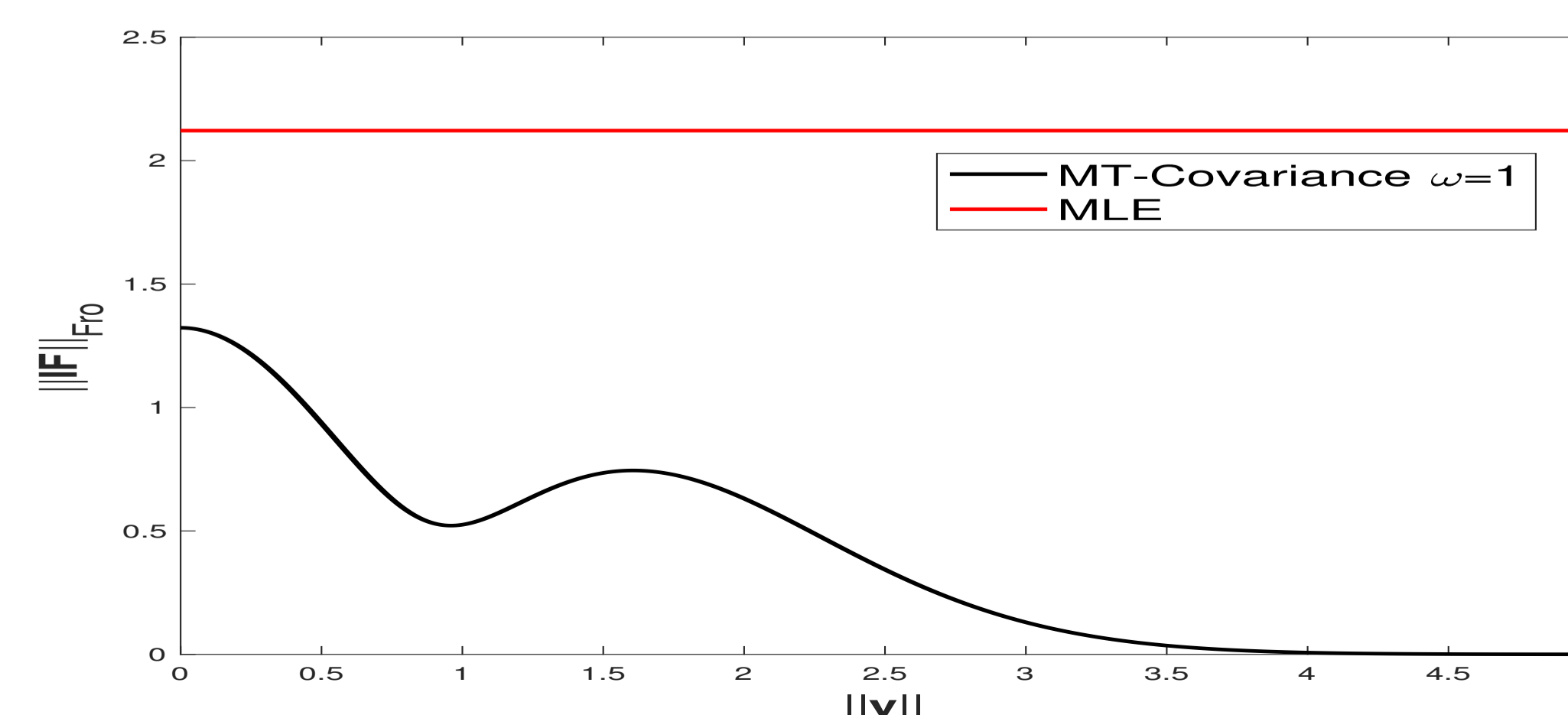
PLUG-IN MEASURE-TRANSFORMED QGLRT

- ▶ Replace $\boldsymbol{\Sigma}_w^{(u)}$ by its empirical estimate obtained from the secondary data:

$$T_u'' \triangleq \frac{\mathbf{a}^H (\hat{\boldsymbol{\Sigma}}_y^{(u)})^{-1} \hat{\mathbf{C}}_x^{(u)} (\hat{\boldsymbol{\Sigma}}_y^{(u)})^{-1} \mathbf{a}}{\mathbf{a}^H (\hat{\boldsymbol{\Sigma}}_y^{(u)})^{-1} \mathbf{a}} \stackrel{H_1}{\leq} t,$$

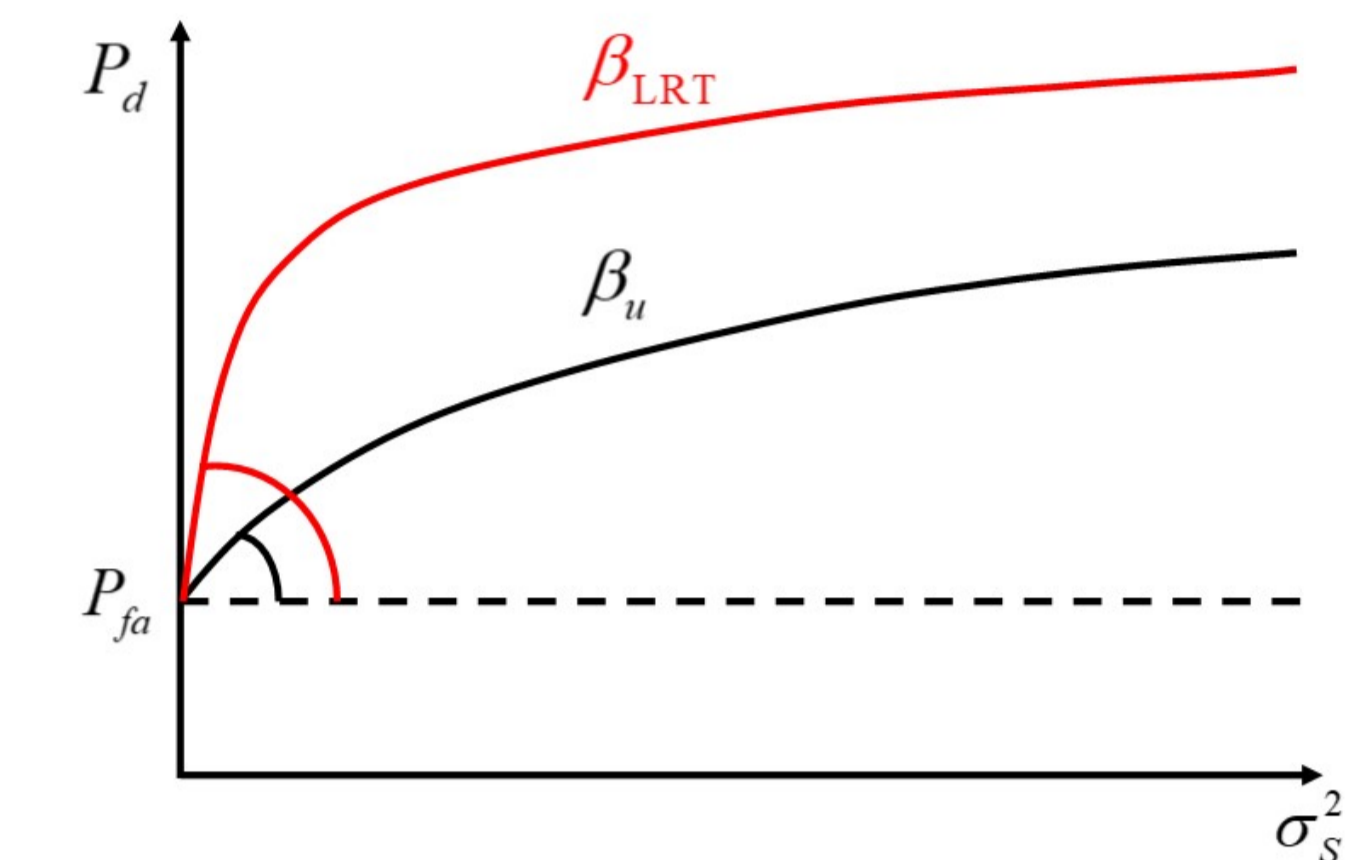
- ▶ Under some mild regularity conditions T_u'' is asymptotically normal.
- ▶ To induce outlier resilience, choose the Gaussian MT-function:

$$u_G(\mathbf{x}; \omega) = \exp\left(-\|\mathbf{P}_a^\perp \mathbf{x}\|^2 / \omega^2\right), \quad \omega \in \mathbb{R}_{++}.$$



SELECTION OF THE GAUSSIAN MT-FUNCTION WIDTH PARAMETER

- ▶ **Principle:** control the asymptotic local power sensitivity to change in the signal variance relative to the omniscient LRT under Gaussian data.



- ▶ **Selection Rule:**

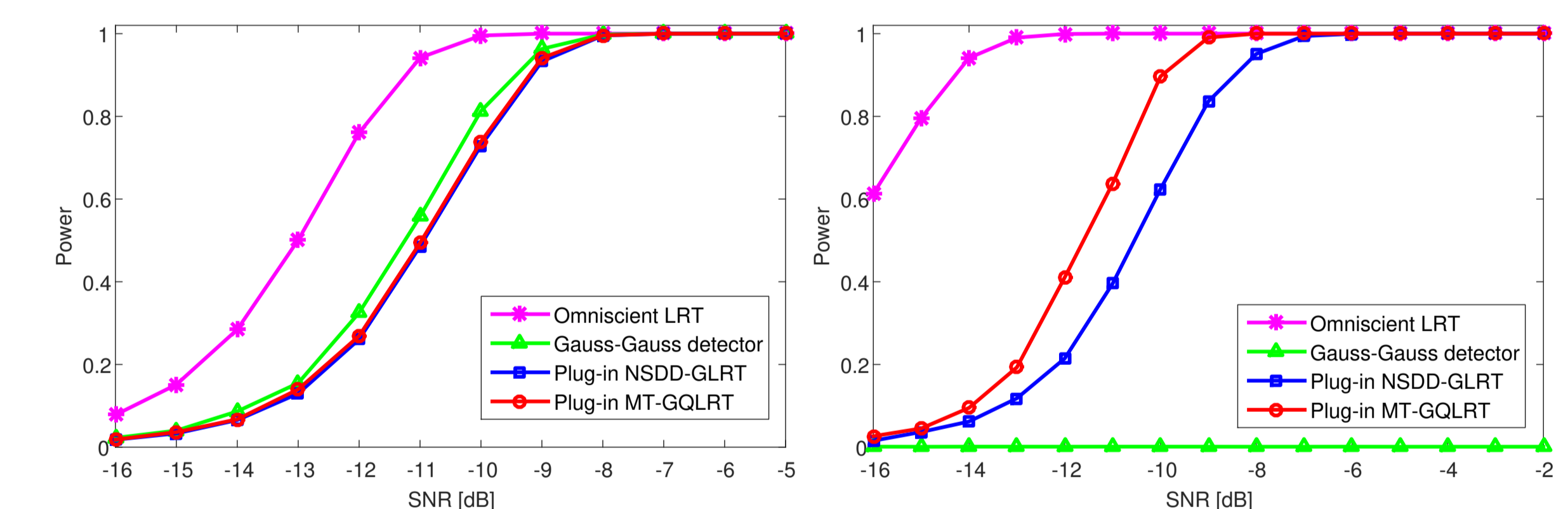
$$\omega_y^* = \inf \left\{ \omega \in \mathbb{R}_+ : \hat{R}_y(\omega) = r \right\}.$$

$$\hat{R}_y(\omega) \triangleq \frac{\partial \hat{\beta}_{u_G}}{\partial \sigma_s^2} / \frac{\partial \hat{\beta}_{LRT}}{\partial \sigma_s^2} \Big|_{\sigma_s^2=0} = \frac{\sqrt{N \max\left(0, \det\left(\mathbf{I}_p - \frac{1}{\omega^2} \left(\mathbf{P}_a^\perp \hat{\boldsymbol{\Sigma}}_y^{(u_G)}(\omega)\right)^2\right)\right)} + Q^{-1}(\alpha)}{\sqrt{N} + Q^{-1}(\alpha)}$$

- ▶ **Constant false alarm rate (CFAR)** w.r.t. the noise power.

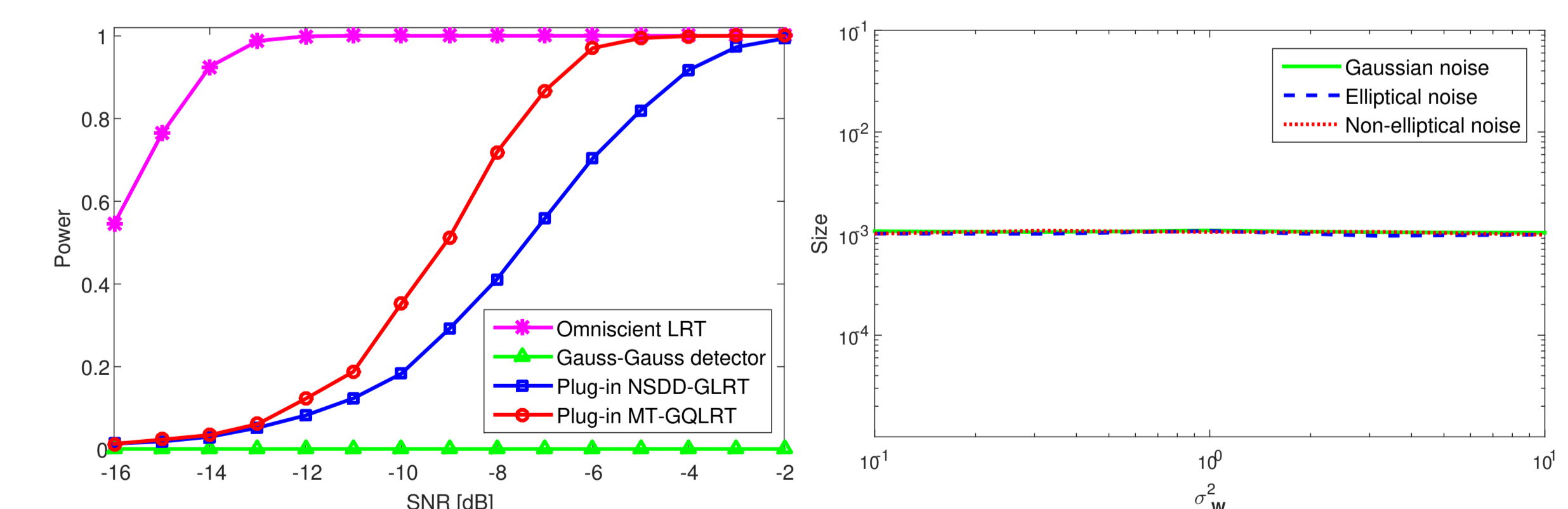
EXAMPLES

- ▶ Parameters: 8-element ULA, $N = M = 500$, $P_{fa} = 10^{-3}$, $r = 0.9$, BPSK signal.



(a) Gaussian noise

(b) Elliptical t -distributed noise



(c) Non-elliptical noise

(d) CFAR analysis

ACKNOWLEDGMENT

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